# Variance Components for Modeling Quantitative Traits

Biostatistics 666
Guest Lecturer
Michael Boehnke

## Today

- Analysis of quantitative traits
- Kinship coefficients
  - measure of genetic similarity between two individuals
- Modeling covariance for pairs of individuals
  - estimating heritability
  - estimating locus-specific heritability
- Extending the model to larger pedigrees

## Kinship Coefficients

 Summarize genetic similarity between pairs of individuals.

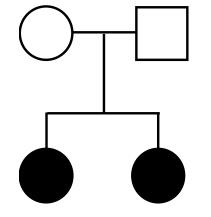
 In a variance components model, they predict the phenotypic similarity between individuals.

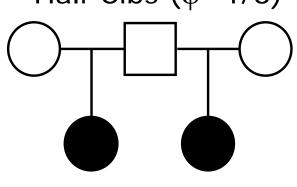
## Kinship Coefficients - Definition

- Given two individuals
  - One with genes (g<sub>i</sub>, g<sub>i</sub>)
  - The other with genes  $(g_k, g_l)$
- The kinship coefficient is:
  - $\frac{1}{4}P(g_i \equiv g_k) + \frac{1}{4}P(g_i \equiv g_l) + \frac{1}{4}P(g_j \equiv g_k) + \frac{1}{4}P(g_j \equiv g_l)$
  - where "≡" represents identity by descent (IBD)
- Probability that two genes sampled at random from each individual are (IBD)

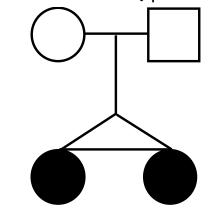
## Some kinship coefficients

Siblings ( $\phi = 1/4$ )





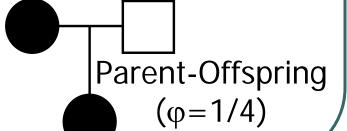
Half-Sibs ( $\phi = 1/8$ ) MZ Twins ( $\phi = 1/2$ )



Unrelated ( $\phi$ =0)







#### What about other relatives?

- For any two related individuals i and j ...
- ... use a recursive algorithm allows calculation of kinship coefficient
- Algorithm requires an order for individuals in the pedigree where ancestors precede descendants
  - That is where for any i>j, i is not ancestor of j
  - Such an order always exists!

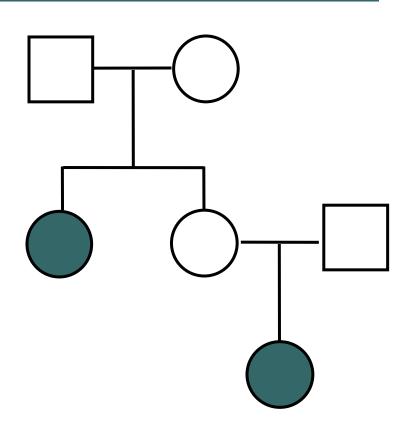
## Computing Kinship Coefficients

The recursive definition is then (for i ≥ j):

$$\varphi_{ij} = \begin{cases} 0 & i \text{ and } j \text{ are founders} \\ \frac{1}{2} & i = j, i \text{ is a founder} \\ \frac{1}{2} (\varphi_{mother(i)j} + \varphi_{father(i)j}) & i \neq j \\ \frac{1}{2} (1 + \varphi_{mother(i)father(i)}) & i = j \end{cases}$$

## An example pedigree...

- Can you find ...
- Suitable ordering for recursive calculation?
- Calculate kinship coefficient between shaded individuals?



#### So far ...

 Summarize genetic similarity between any two individuals ...

 Next, we will proceed to build a simple model for their phenotypes

## Simplest Data Structure

- Pairs of related individuals
  - Siblings (or twins!)
  - Parent-Offspring
- Corresponding phenotype measurements
  - $y = (y_1, y_2)$

## Elements for a simple model ...

- If the trait is normally distributed ...
- Model mean and variance for y<sub>1</sub> and y<sub>2</sub>
  - Mean and variance could be assumed equal ...
  - ... or they could depend on some covariates
- But we are also interested in covariance between the two ...

#### Variance-Covariance Matrix

$$\Omega = \begin{bmatrix} V(y_1) & Cov(y_1, y_2) \\ Cov(y_1, y_2) & V(y_2) \end{bmatrix}$$

Model must describe not only variance of each observation but also covariance for pairs of observations

## Bivariate density function

Normal density function

$$L(y) = \frac{1}{\sqrt{2\pi}} \sigma^{-1} e^{-\frac{1}{2}(y-\mu)^2/\sigma^2}$$

Bivariate normal density function

$$L(\mathbf{y}) = \frac{1}{2\pi} |\Omega|^{-1/2} e^{-\frac{1}{2}(\mathbf{y} - \mathbf{\mu})'\Omega^{-1}(\mathbf{y} - \mathbf{\mu})}$$

Extends univariate density function

#### Intuition on Normal Densities

$$L(y) = (2\pi)^{-\frac{1}{2}} \sigma^{-1} e^{-\frac{1}{2}(y-\mu)^2/\sigma^2}$$

Scaling parameter, penalizes settings with large variances

Distance between observation and its expected value

#### Bivariate Normal Densities

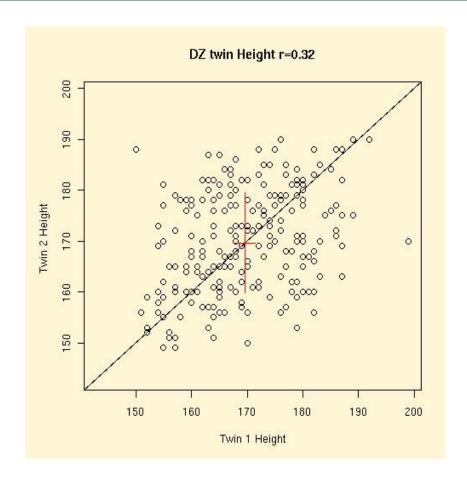
$$L(\mathbf{y}) = (2\pi)^{-1} \left( |\Omega|^{-\frac{1}{2}} e^{-\frac{1}{2}(\mathbf{y} - \boldsymbol{\mu})'\Omega^{-1}(\mathbf{y} - \boldsymbol{\mu})} \right)$$

Scaling parameter, penalizes settings with large variances

Distance between observation and its expected value

## Possible Application...

In a sample of twin or sibling pairs, we could use all the data to estimate means, variances and even covariances...



(Data from David Duffy)<sub>6</sub>

#### Incorporating Kinship Coefficients

- If genes influence trait ...
- Covariance will differ for each class of relative pair
- Instead of estimating covariance for each relationship, ...
- Impose genetic model that incorporates kinship and relates covariance between different classes of relative pair

# A Simple Model for the Variance-Covariance Matrix

$$\Omega = egin{bmatrix} \sigma_g^2 + \sigma_e^2 & 2 arphi \sigma_g^2 \ 2 arphi \sigma_g^2 & \sigma_g^2 + \sigma_e^2 \end{bmatrix}$$

Where,

 $\varphi$  is the kinship coefficient for the two individuals

# Example...

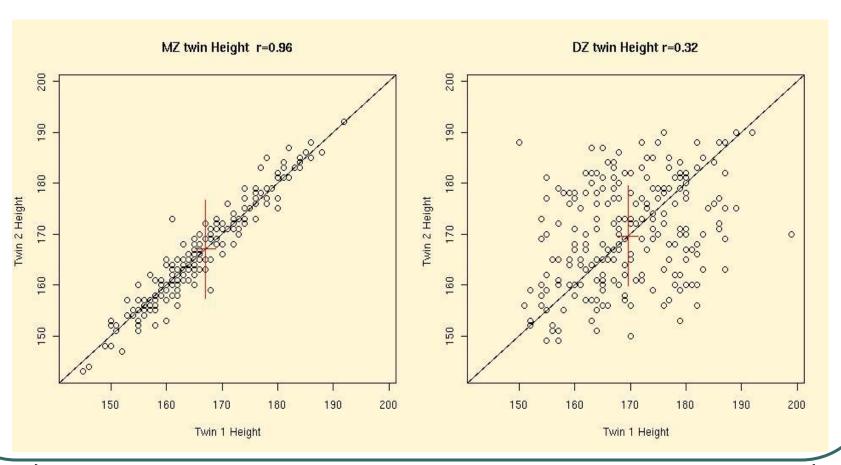
	N	r
MZ males	292	.80
MZ females	380	.80
DZ males	179	.47
DZ females	184	.55
DZ male-female	284	.41

(Reading ability scores from Eaves et al., 1997)

### Interpretation...

- Fitting a maximum likelihood model...
  - Eaves et. al estimated
    - $\sigma_{\rm g}^{2} = .81$
    - $^{\bullet}$   $\sigma_{\rm e}^{2} = .19$
  - Found no evidence for sex differences
  - Saturated model did not improve fit

## Height in DZ and MZ twins



(How would you interpret these data from David Duffy?)

#### So far ...

 Model allows us to estimate the genetic contribution to the variation in any trait

Incorporates different relative pairs ...

- But it doesn't always fit...
  - Fortunately, the model can be easily refined

## Another Example...

	N	r
MZ males	271	.56
MZ females	353	.52
DZ males	167	.33
DZ females	165	.45
DZ male-female	260	.41

(Psychomotor retardation scores from Eaves et al., 1997)

#### Refined Matrix

$$\Omega = \begin{bmatrix} \sigma_g^2 + \sigma_c^2 + \sigma_e^2 & 2\varphi\sigma_g^2 + \sigma_c^2 \\ 2\varphi\sigma_g^2 + \sigma_c^2 & \sigma_g^2 + \sigma_c^2 + \sigma_e^2 \end{bmatrix}$$

Where,

 $\varphi$  is the kinship coefficient for the two individuals

### Interpretation...

- Fitting a maximum likelihood model...
  - Eaves et. al estimated (for males)

• 
$$\sigma_{\rm g}^{2} = .29$$

• 
$$\sigma_{c}^{2} = .24$$

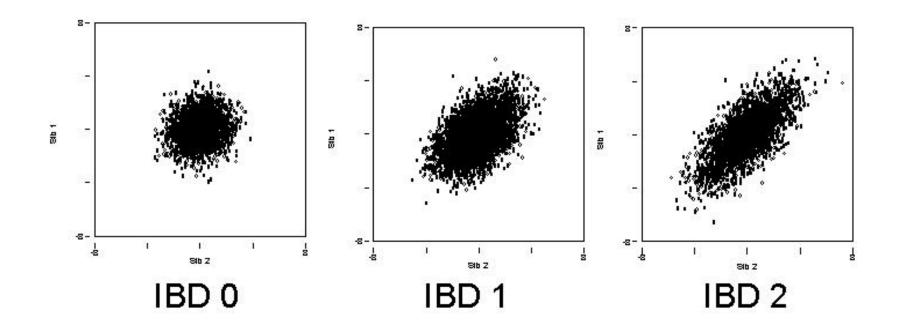
• 
$$\sigma_e^2 = .46$$

• Additive genetic effects could not explain similarities. Any idea why?

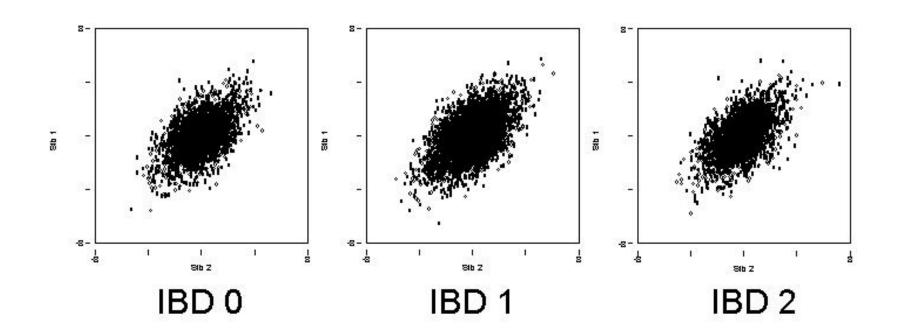
#### Incorporating IBD Coefficients

- Covariance might differ according to sharing at a particular locus
  - If locus contains genes that influence the trait
- Again, impose a genetic model and estimate model parameters

# Linkage



# No Linkage



#### Relationship to IBD probabilities

 For non-inbred pair of relatives, marker or locus-specific kinship coefficients can be derived from IBD probabilities:

$$\varphi_{marker} = \frac{1}{4}P(IBD_{marker} = 1) + \frac{1}{2}P(IBD_{marker} = 2)$$

#### Variance-Covariance Matrix

$$\Omega = \begin{bmatrix} \sigma_a^2 + \sigma_g^2 + \sigma_e^2 & 2\varphi_{marker}\sigma_a^2 + 2\varphi\sigma_g^2 \\ 2\varphi_{marker}\sigma_a^2 + 2\varphi\sigma_g^2 & \sigma_a^2 + \sigma_g^2 + \sigma_e^2 \end{bmatrix}$$

Where,

 $\varphi$  is the kinship coefficient for the two individuals  $\varphi_{marker}$  depends on the number of alleles shared IBD

# Likelihood function, Incorporating Uncertain IBD

$$L = \prod_{i} \sum_{j=0,1,2} Z_{ij} (2\pi)^{-1} |\Omega_{IBD=j}|^{-\frac{1}{2}} e^{-\frac{1}{2}(\mathbf{y}-\mathbf{\mu})'\mathbf{\Omega}_{IBD=j}^{-1}(\mathbf{y}-\mathbf{\mu})}$$

$$\approx \prod_{i} (2\pi)^{-1} |\Omega^*|^{-\frac{1}{2}} e^{-\frac{1}{2}(\mathbf{y}-\mathbf{\mu})'\mathbf{\Omega}^{*-1}(\mathbf{y}-\mathbf{\mu})}$$

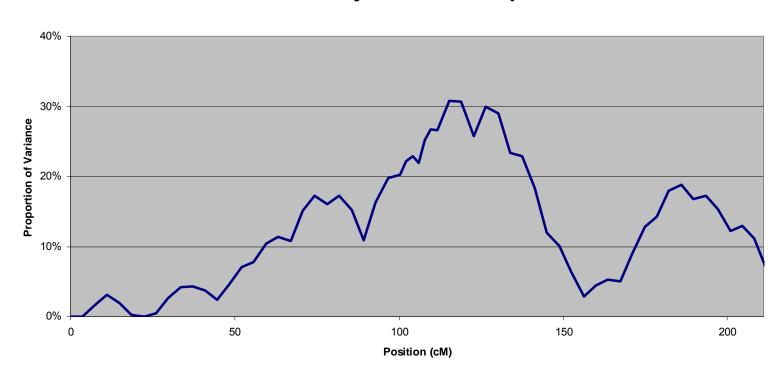
$$Z_{ij} = P(IBD_i = j \mid \text{marker data})$$
 IBD sharing probabilities 
$$\Omega^* = \sum Z_{ij} \Omega_{IBD=j}$$
 "Expected"  $\Omega$ 

#### How it works ...

- To find linkage to a particular trait...
- Collect sibling pair sample
- Calculate IBD for multiple points along genome
- Model covariance as a function of IBD sharing at each point

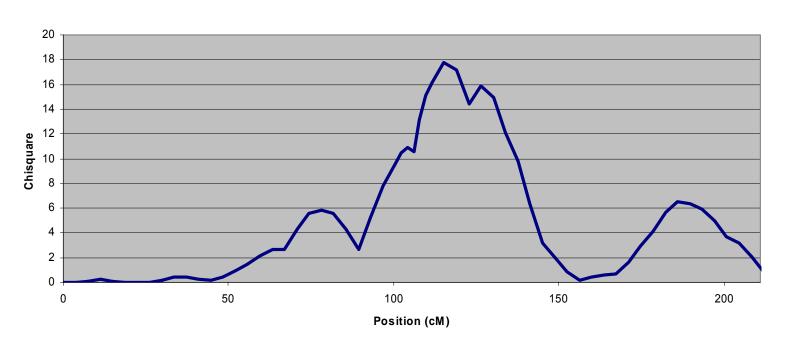
# Example...

#### **Estimated Major Gene Component**



# Example...

#### **Likelihood Ratio Chisquare**



#### So far ...

Models for similarity between relative pairs

Kinship coefficient used to estimate overall genetic effect

 Locus-specific coefficients used to detect genetic linkage

#### Extensions ...

The model extends gracefully to other settings:

For larger pedigrees, we extend the covariance matrix

 To model genetic association, we model specific means for each individual

## Larger Pedigrees...

$$\Omega_{jk} = \begin{cases} \sigma_a^2 + \sigma_g^2 + \sigma_e^2 & \text{if } j = k \\ 2\varphi_{marker}\sigma_a^2 + 2\varphi\sigma_g^2 & \text{if } j \neq k \end{cases}$$

Where,

 $\varphi$  is the kinship coefficient for the two individuals  $\varphi_{marker}$  depends on the number of alleles shared IBD j and k index different individuals in the family

## Multivariate density function

Normal density function

$$L(y) = (2\pi)^{-\frac{1}{2}} \sigma^{-1} e^{-\frac{1}{2}(y-\mu)^2/\sigma^2}$$

Multivariate normal density function

$$L(\mathbf{y}) = 2\pi^{-n/2} |\Omega|^{-1/2} e^{-1/2(\mathbf{y} - \mathbf{\mu})'\Omega^{-1}(\mathbf{y} - \mathbf{\mu})}$$

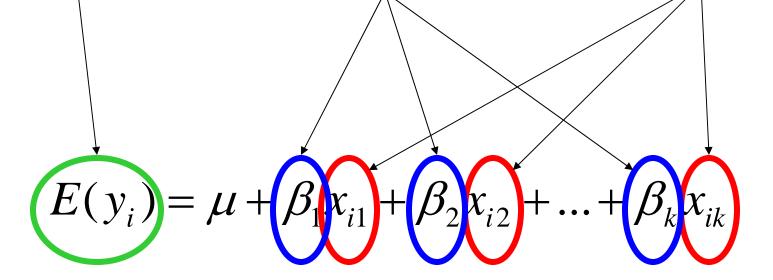
Extends univariate density function

#### Means Model

Expected Phenotype for Individual *i* (e.g. expected weight)

Estimated effects for covariates (e.g. expected weight increases 1kg/year with age)

Measured Covariates for Individual *i* (e.g. age, sex, ...)



In addition to modeling variances and covariances, we can model the means

## Simple Association Model

- Each copy of allele changes trait by a fixed amount
  - Include covariate counting copies for allele of interest
- Evidence for association when a ≠ 0

$$E(y_i) = \mu + a * [number of copies of mutant allele]$$
  
 $E(y_i) = \mu + \beta_X X_i$ 

X is the number of copies for allele of interest.  $\beta_x$  is the estimated effect of each copy (the additive genetic value).

## Today

- Analysis of quantitative traits
- Kinship coefficients
  - Measure of genetic similarity between two individuals
- Modeling covariance for pairs of individuals
  - estimating heritability
  - estimating locus-specific heritability
- Extending the model to larger pedigrees

#### Useful References

- Amos (1994)
   Am J Hum Genet **54**:535-543
- Hopper and Matthews (1982)
   Ann Hum Genet 46:373–383
- Lange and Boehnke (1983)
   Am J Med Genet 14:513-24