

A Simple Gibbs Sampler

Biostatistics 615/815

Lecture 20

Scheduling

- No Class
 - November 30
- Review session
 - Thursday , December 7
- Final Assessment
 - Tuesday, December 12

Optimization Strategies

- **Single Variable**
 - Golden Search
 - Quadratic Approximations

- **Multiple Variables**
 - Simplex Method
 - E-M Algorithm
 - Simulated Annealing

Simulated Annealing

- Stochastic Method
- Sometimes takes up-hill steps
 - Avoids local minima
- Solution is gradually *frozen*
 - Values of parameters with largest impact on function values are fixed earlier

Gibbs Sampler

- Another MCMC Method
- Update a single parameter at a time
- Sample from conditional distribution when other parameters are fixed

Gibbs Sampler Algorithm

Consider a particular choice of parameter values $\boldsymbol{\theta}^{(t)}$

Define the next set of parameter values by :

a. Selecting component to update, say i

b. Sample value for $\theta_i^{(t+1)}$ from $p(\theta_i | x, \theta_1, \theta_2, \dots, \theta_{i-1}, \theta_{i+1}, \dots, \theta_k)$

Increment t and repeat previous steps.

Alternative Algorithm

Consider a particular choice of parameter values $\theta^{(t)}$

Define the next set of parameter values by :

- a. Update each component, $1 \dots k$, in turn
- b. Sample value for $\theta_1^{(t+1)}$ from $p(\theta_1 | x, \theta_2, \theta_3, \dots, \theta_k)$
- c. Sample value for $\theta_2^{(t+1)}$ from $p(\theta_2 | x, \theta_1, \theta_3, \dots, \theta_k)$
- ...
- z. Sample value for $\theta_k^{(t+1)}$ from $p(\theta_k | x, \theta_1, \theta_3, \dots, \theta_{k-1})$

Increment t and repeat previous steps.

Key Property: Stationary Distribution

Suppose that $(\theta_1^{(t)}, \theta_2^{(t)}, \dots, \theta_k^{(t)}) \sim p(\theta_1, \theta_2, \dots, \theta_k | x)$

Then $(\theta_1^{(t+1)}, \theta_2^{(t)}, \dots, \theta_k^{(t)})$ is distributed as

$$p(\theta_1 | \theta_2, \dots, \theta_k, x) p(\theta_2, \dots, \theta_k | x) = p(\theta_1, \theta_2, \dots, \theta_k | x)$$

In fact...

$$\boldsymbol{\theta}^{(t)} \sim p(\boldsymbol{\theta} | x) \Rightarrow \boldsymbol{\theta}^{(t+1)} \sim p(\boldsymbol{\theta} | x)$$

Eventually, we expect the Gibbs sampler to sample parameter values from their posterior distribution

Gibbs Sampling for Mixture Distributions

- Sample each of the mixture parameters from conditional distribution
 - Dirichlet, Normal and Gamma distributions are typical
- Simple alternative is to sample the *source* of each observation
 - Assign observation to specific component

Sampling A Component

$$\Pr(Z_j = i \mid x_j, \boldsymbol{\pi}, \boldsymbol{\phi}, \boldsymbol{\eta}) = \frac{\pi_i f(x_j \mid \phi_i, \boldsymbol{\eta})}{\sum_l \pi_l f(x_j \mid \phi_l, \boldsymbol{\eta})}$$

- Calculate the probability that the observation originated from a specific component...
- ... can you recall how random numbers can be used to sample from one of a few discrete categories?

C Code: Sampling A Component

```
int sample_group(double x, int k,
                 double * probs, double * mean, double * sigma)
{
    int group; double p = Random();
    double lk = dmix(x, k, probs, mean, sigma);

    for (group = 0; group < k - 1; group++)
    {
        double pgroup = probs[group] *
                        dnorm(x, mean[group], sigma[group])/lk;

        if (p < pgroup)
            return group;

        p -= pgroup;
    }

    return k - 1;
}
```

Calculating Mixture Parameters

$$n_i = \sum_{j:Z_j=i} 1$$

$$p_i = n_i / n$$

$$\bar{x}_i = \sum_{j:Z_j=i} x_j / n_i$$

$$s_i = \sqrt{\left(\sum_{j:Z_j=i} x_j^2 - n_i \bar{x}_i^2 \right) / n_i}$$

- Before sampling a new origin for an observation...
- ... update mixture parameters given current assignments
- Could be expensive!

C Code: Updating Mixture Parameters II

```
void remove_observation(double x, int group,  
                       double * counts, double * sum, double * sumsq)  
{  
    counts[group] --;  
    sum[group] -= x;  
    sumsq[group] -= x * x;  
}
```

```
void add_observation(double x, int group,  
                   double * counts, double * sum, double * sumsq)  
{  
    counts[group] ++;  
    sum[group] += x;  
    sumsq[group] += x * x;  
}
```

Selecting a Starting State

- Must start with an assignment of observations to groupings
- Many alternatives are possible, I chose to perform random assignments with equal probabilities...

C Code: Starting State

```
void initial_state(int k, int * group,
                  double * counts, double * sum, double * sumsq)
{
    int i;

    for (i = 0; i < k; i++)
        counts[i] = sum[i] = sumsq[i] = 0.0;

    for (i = 0; i < n; i++)
    {
        group[i] = Random() * k;

        counts[group[i]] ++;
        sum[group[i]] += data[i];
        sumsq[group[i]] += data[i] * data[i];
    }
}
```


The Gibbs Sampler

- Select initial state
- Repeat a large number of times:
 - Select an element
 - Update conditional on other elements
- If appropriate, output summary for each run...

C Code: Core of The Gibbs Sampler

```
initial_state(k, probs, mean, sigma, group, counts, sum, sumsq);
for (i = 0; i < 10000000; i++)
{
    int id = Random() * n;

    if (counts[group[id]] < MIN_GROUP) continue;

    remove_observation(data[id], group[id], counts, sum, sumsq);
    update_estimates(k, n - 1, probs, mean, sigma,
                    counts, sum, sumsq);
    group[id] = sample_group(data[id], k, probs, mean, sigma);
    add_observation(data[id], group[id], counts, sum, sumsq);

    if ((i > BURN_IN) && (i % THIN_INTERVAL == 0))
        /* Collect statistics */
}
```

Gibbs Sampler: Memory Allocation and Freeing

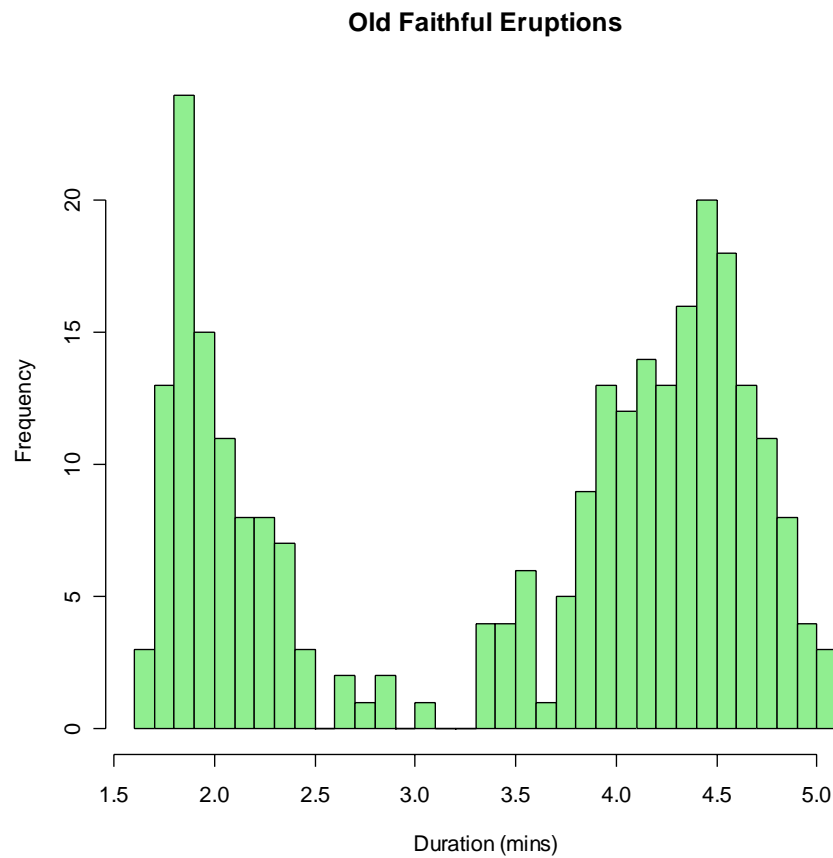
```
void gibbs(int k, double * probs, double * mean, double * sigma)
{
    int i, * group = (int *) malloc(sizeof(int) * n);
    double * sum = alloc_vector(k);
    double * sumsq = alloc_vector(k);
    double * counts = alloc_vector(k);

    /* Core of the Gibbs Sampler goes here */

    free_vector(sum, k);
    free_vector(sumsq, k);
    free_vector(counts, k);
    free(group);
}
```

Example Application

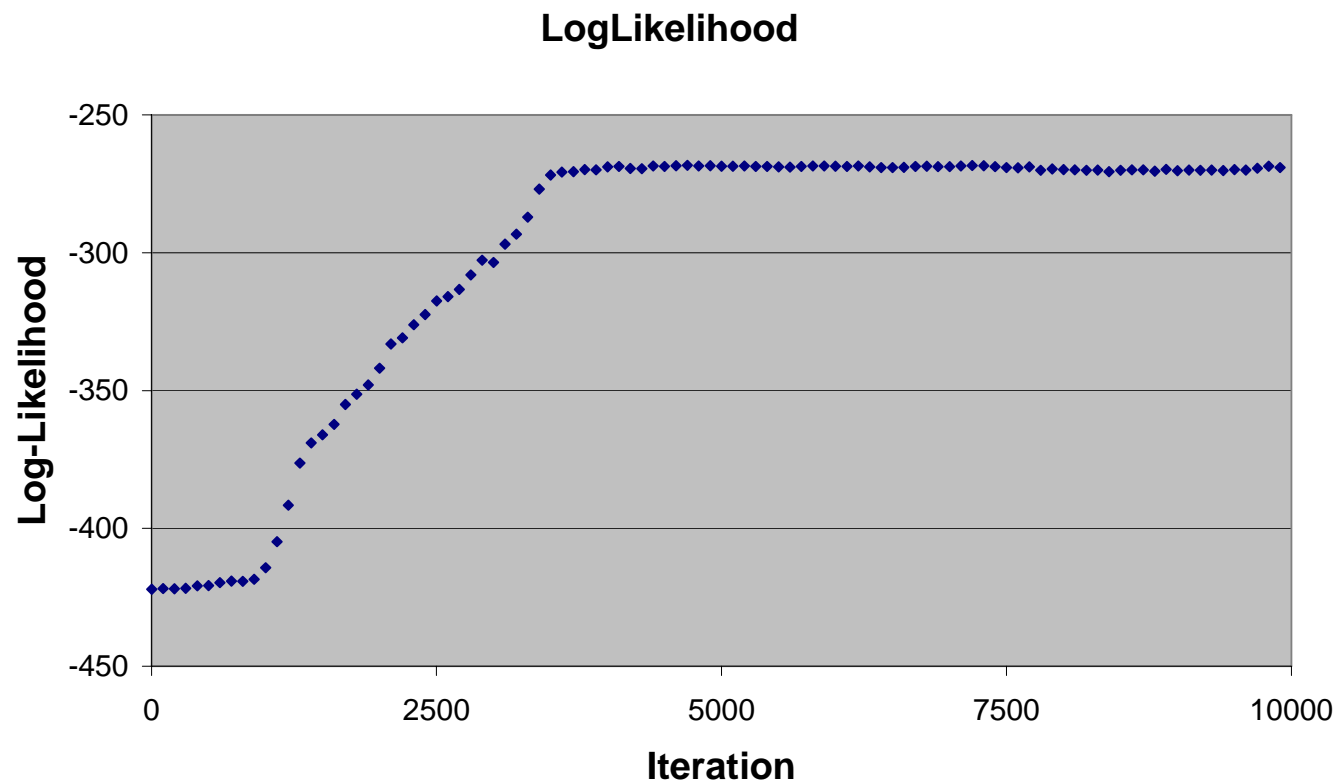
Old Faithful Eruptions (n = 272)



Notes ...

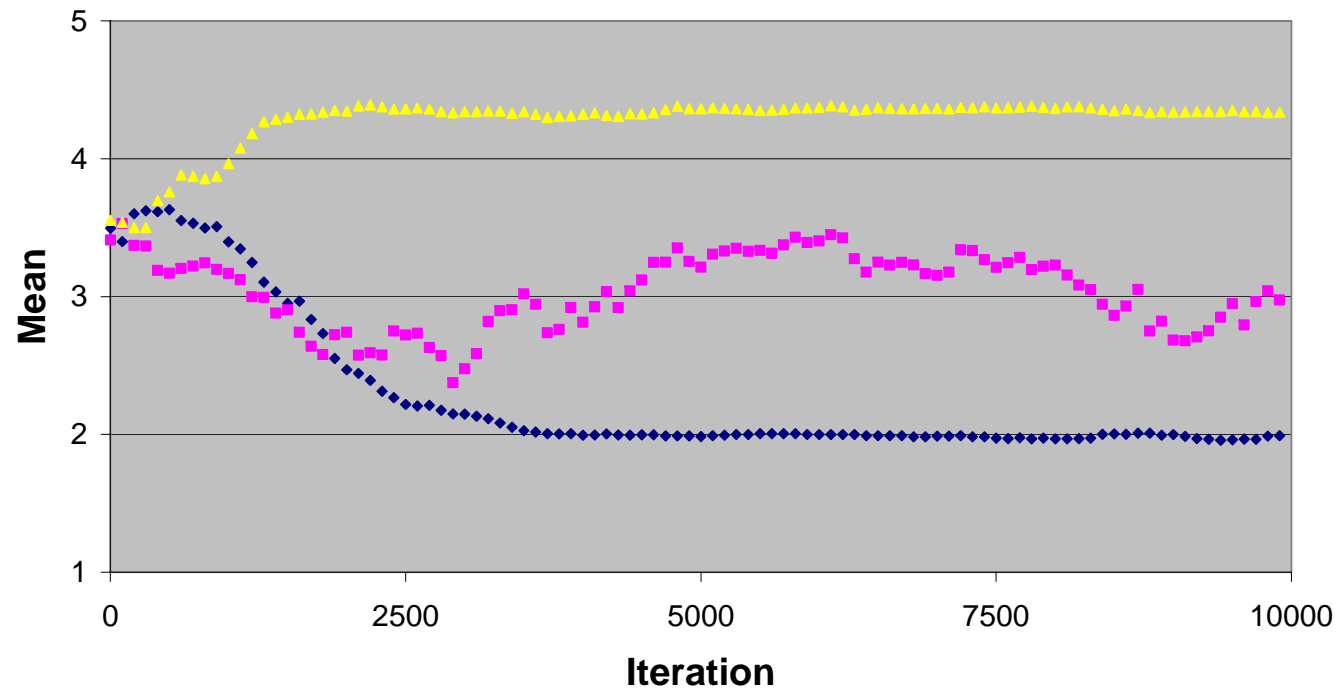
- My first few runs found excellent solutions by fitting components accounting for very few observations but with variance near 0
- Why?
 - Repeated values due to rounding
 - To avoid this, I set MIN_GROUP to 13 (5%)

Gibbs Sampler Burn-In

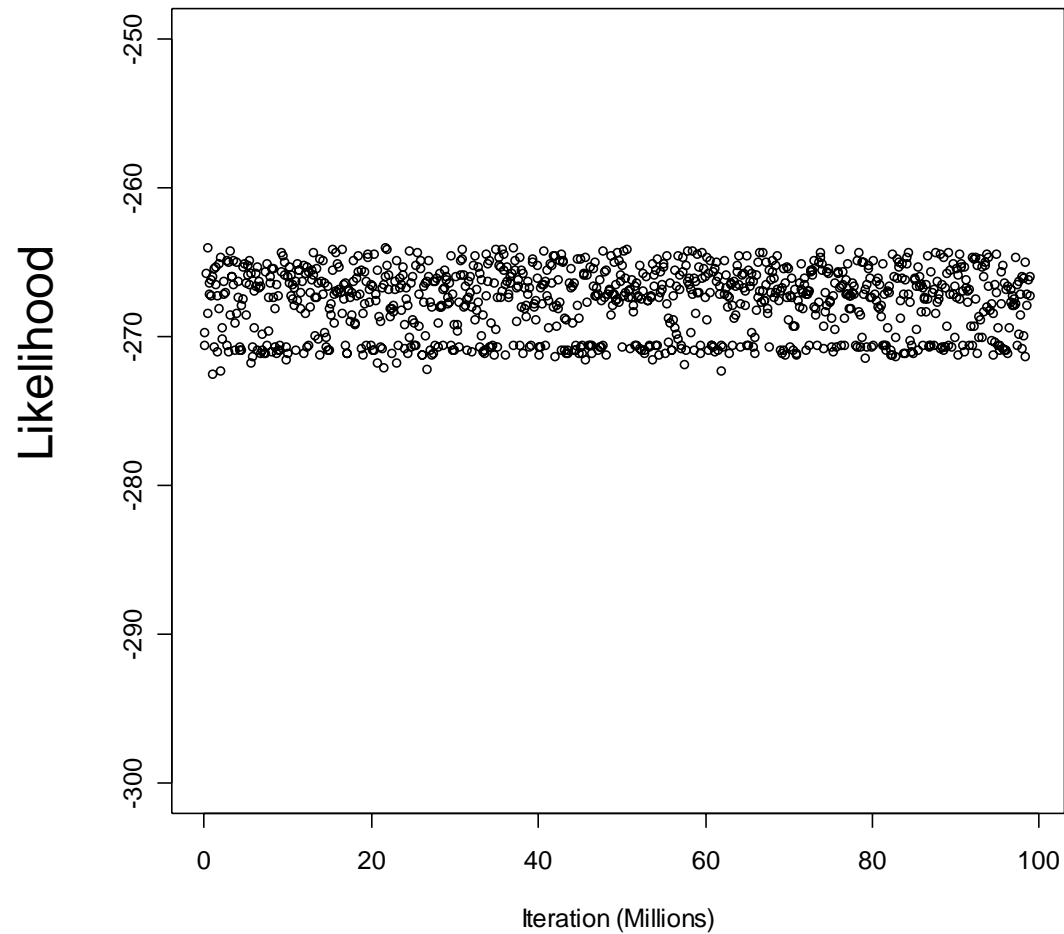


Gibbs Sampler Burn-In

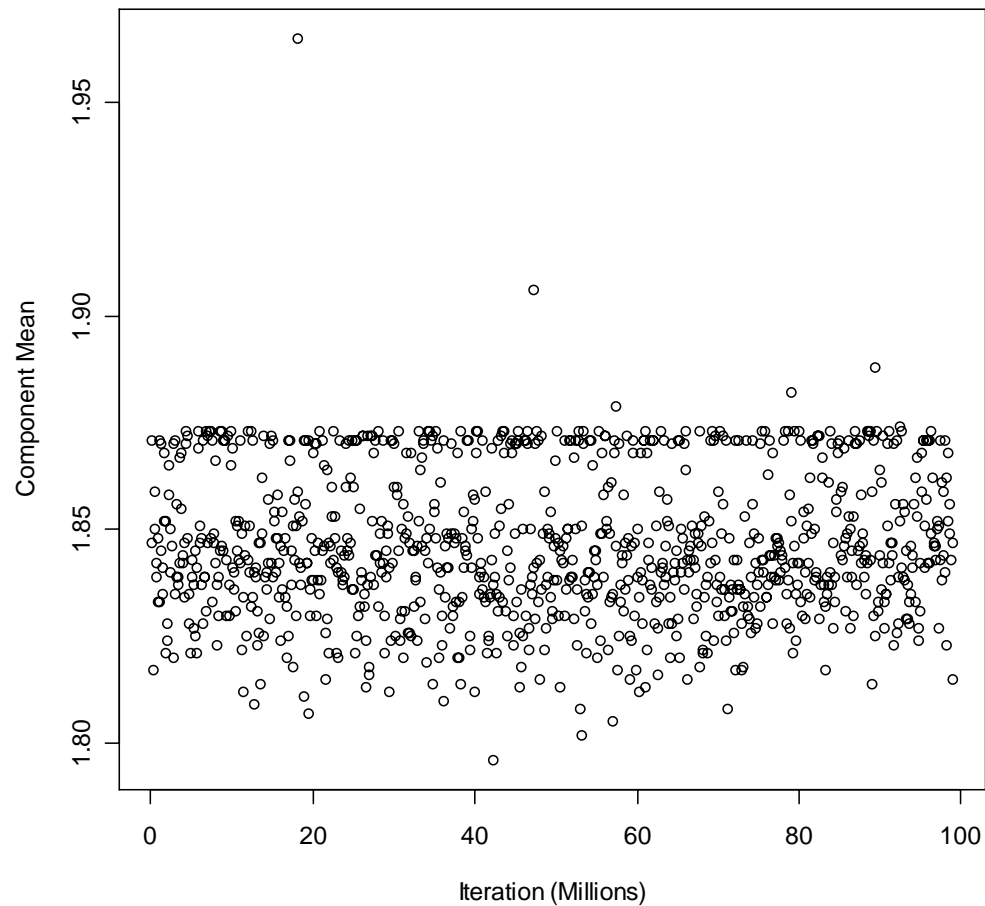
Mixture Means



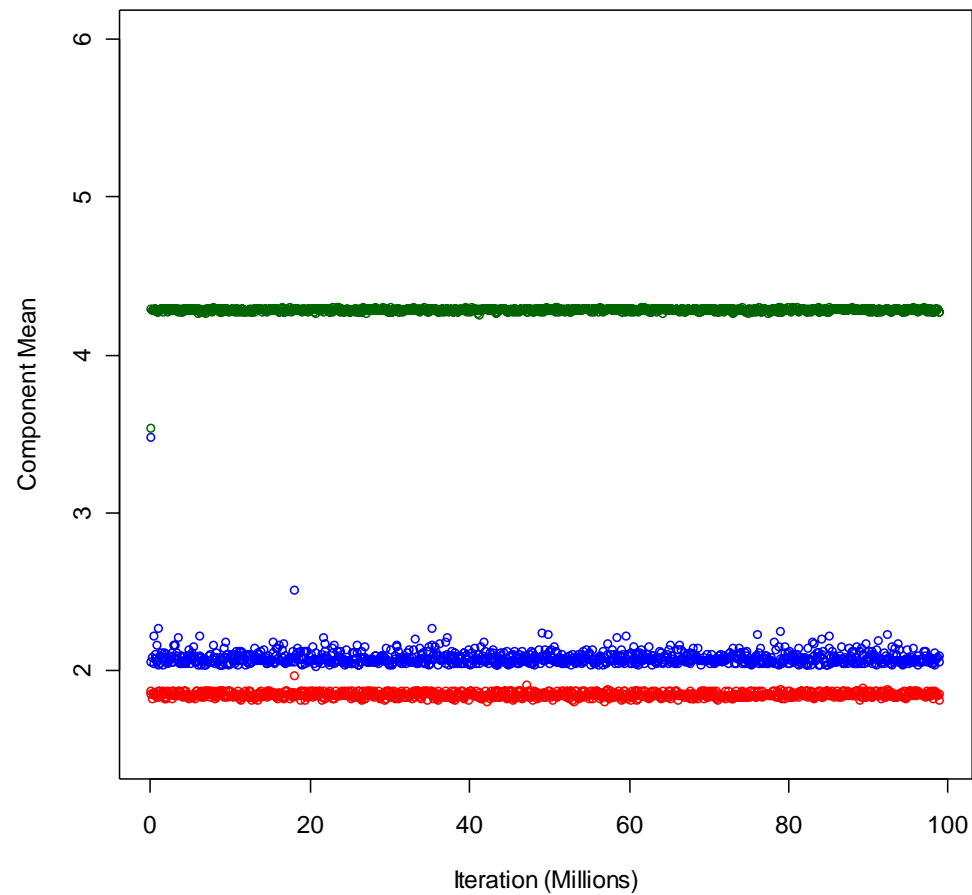
Gibbs Sampler After Burn-In Likelihood



Gibbs Sampler After Burn-In Mean for First Component



Gibbs Sampler After Burn-In



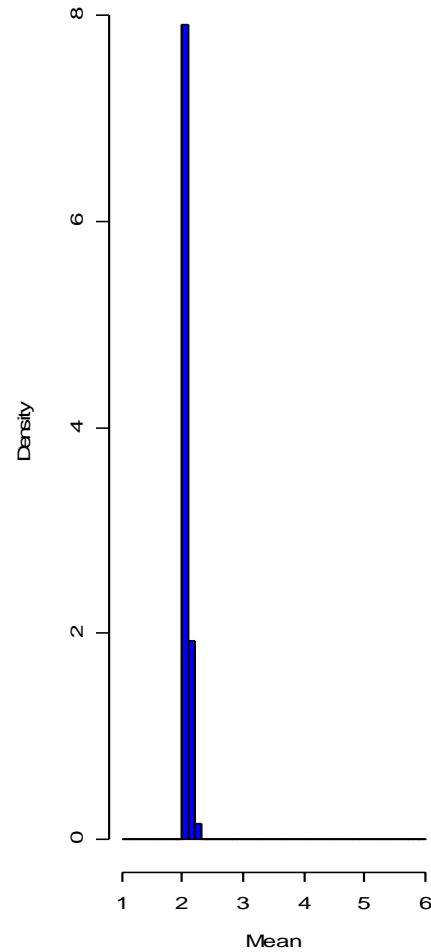
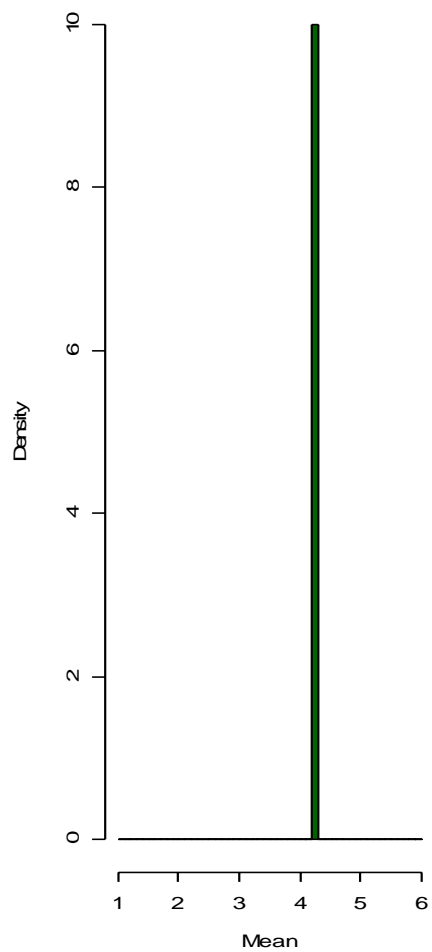
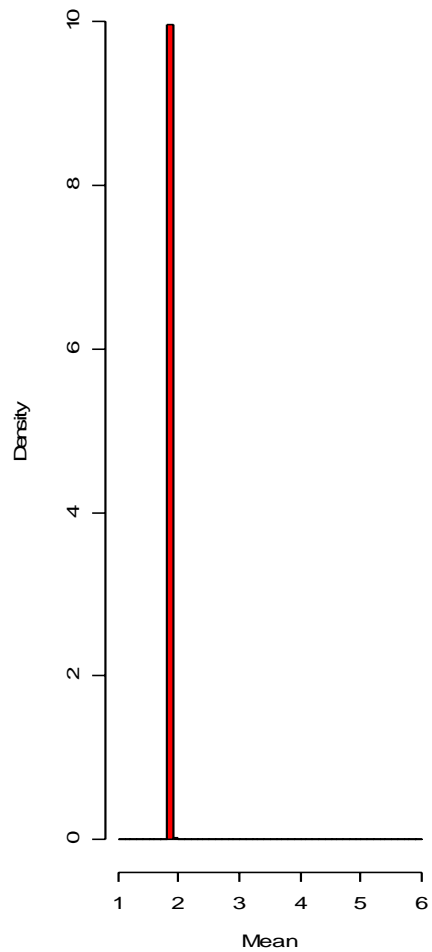
Notes on Gibbs Sampler

- Previous optimizers settled on a minimum eventually
- The Gibbs sampler continues wandering through the stationary distribution...
- Forever!

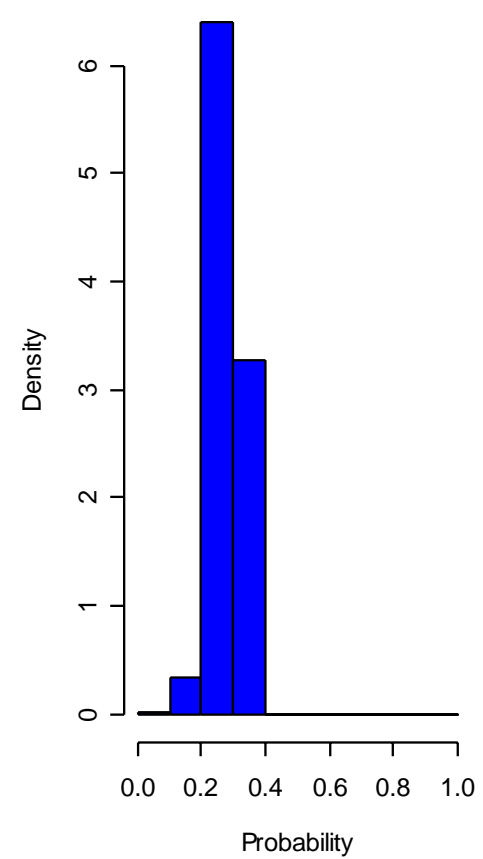
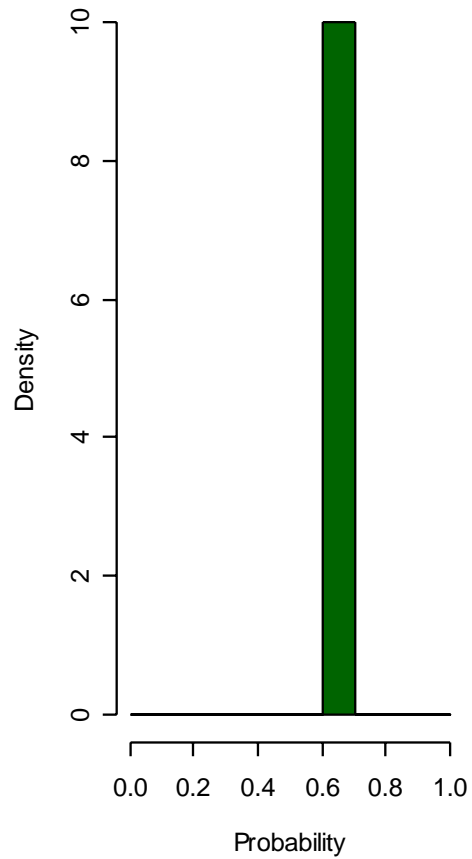
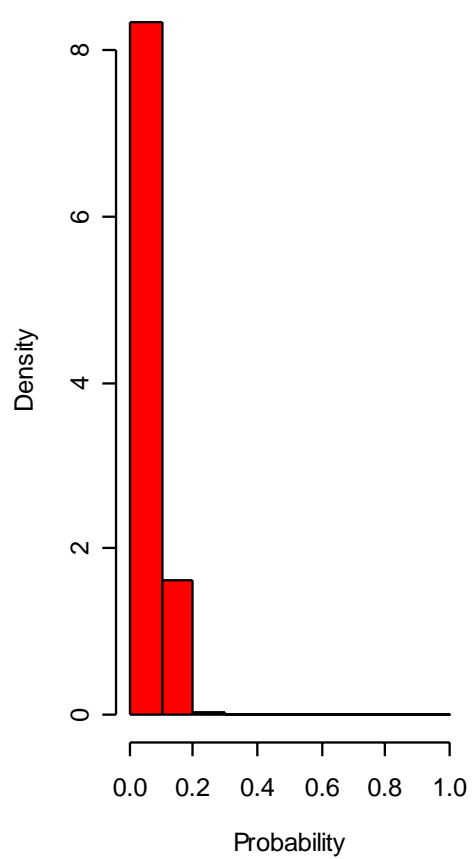
Drawing Inferences...

- To draw inferences, summarize parameter values from stationary distribution
- For example, might calculate the mean, median, etc.

Component Means



Component Probabilities



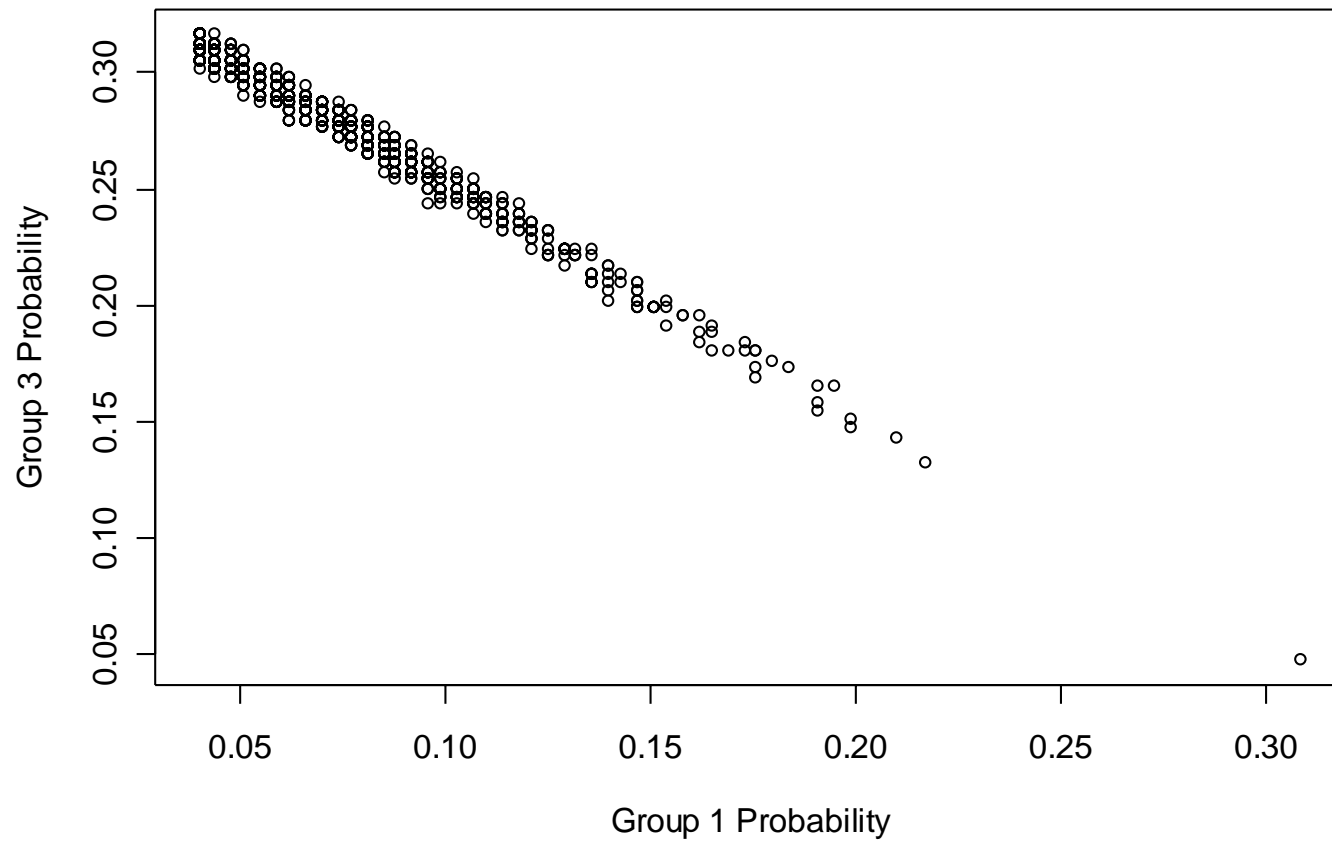
Overall Parameter Estimates

- The means of the posterior distributions for the three components were:
 - Frequencies of 0.073, 0.278 and 0.648
 - Means of 1.85, 2.08 and 4.28
 - Variances of 0.001, 0.065 and 0.182
- Our previous estimates were:
 - Components contributing .160, 0.195 and 0.644
 - Component means are 1.856, 2.182 and 4.289
 - Variances are 0.00766, 0.0709 and 0.172

Joint Distributions

- Gibbs Sampler provides other interesting information and insights
- For example, we can evaluate joint distribution of two parameters...

Component Probabilites



So far today ...

- Introduction to Gibbs sampling
- Generating posterior distributions of parameters conditional on data
- Providing insight into joint distributions

A little bit of theory

- Highlight connection between Simulated Annealing and the Gibbs sampler ...
- Fill in some details of the Metropolis algorithm

Both Methods Are Markov Chains

- The probability of any state being chosen depends only on the previous state

$$\Pr(S_n = i_n \mid S_{n-1} = i_{n-1}, \dots, S_0 = i_0) = \Pr(S_n = i_n \mid S_{n-1} = i_{n-1})$$

- States are updated according to transition matrix with elements p_{ij} . This matrix defines important properties, including **periodicity** and **irreducibility**.

Metropolis-Hastings Acceptance Probability

Let $q_{ij} = q(\text{propose } S_{n+1} = j \mid S_n = i)$

Let π_i and π_j be the relative probabilities of each state

The Metropolis - Hastings acceptance probability is :

$$a_{ij} = \min\left(1, \frac{\pi_j q_{ji}}{\pi_i q_{ij}}\right) \quad \text{or} \quad a_{ij} = \min\left(1, \frac{\pi_j}{\pi_i}\right) \text{ if } q_{ij} = q_{ji}$$

Only the ratio $\frac{\pi_j}{\pi_i}$ must be known, not the actual values of π

Metropolis-Hastings Equilibrium

If we use the Metropolis - Hastings algorithm to update a Markov Chain, it will reach an equilibrium distribution where $\Pr(S = i) = \pi_i$

For this to happen, the proposal density must allow all states to communicate.

Gibbs Sampler

The Gibbs sampler ensures that $\pi_i q_{ij} = \pi_j q_{ji}$

As a consequence, $a_{ij} = \min\left(1, \frac{\pi_j q_{ji}}{\pi_i q_{ij}}\right) = 1$

Simulated Annealing

Given a temperature parameter τ ,

replace π_i with $\pi_i^{(\tau)} = \frac{\pi_i^{\frac{1}{\tau}}}{\sum_j \pi_j^{\frac{1}{\tau}}}$

At high temperatures, the probability distribution is flattened

At low temperatures, larger weights are given to high probability states

Additional Reading

- If you need a refresher on Gibbs sampling
 - Bayesian Methods for Mixture Distributions
M. Stephens (1997)
<http://www.stat.washington.edu/stephens/>
- Numerical Analysis for Statisticians
 - Kenneth Lange (1999)
 - Chapter 24