## Replacing IBS with IBD: The MLS Method

Biostatistics 666
Lecture 15

## Previous Lecture

- Analysis of Affected Relative Pairs
- Test for Increased Sharing at Marker
- Expected Amount of IBS Sharing


## Previous Lecture: Expected IBS Sharing

- Calculated probability of IBS for each IBD state
- Probability of IBD state depends on relationship
- Under the null hypothesis of no linkage

$$
P(I B S=i \mid R)=\sum_{j=0}^{2} P(I B S=i \mid I B D=j) P(I B D=j \mid R)
$$

## Shortcomings of IBS Method

- All sharing is weighted equally
- Sharing a rare allele
- Sharing a common allele
- Sharing homozygous genotype
- Sharing heterozygous genotype
- Inefficient.
- Data contains additional information that is being ignored.


## Today

- A likelihood based approach
- Evaluate linkage in fully informative pairs
- An E-M algorithm for practical settings
- MLS method, Risch (1990)


## Simple Case

- If IBD could be observed
- Each pair of individuals scored as
- IBD=0
- $1 B D=1$
- $\operatorname{IBD}=2$
- Evaluate likelihood for null and alternative hypothesis


## The Model

- Depends on three parameters $\mathrm{z}_{0}, \mathrm{z}_{1}, \mathrm{z}_{2}$
${ }^{\bullet}$ Probability of sharing 0,1 and 2 alleles IBD
- Under the null, determined by relationship
- Under the alternative, determined by genetic model


## Sib Pair Likelihood (Fully Informative Data)

Under the null hypothesis:

$$
L=(1 / 4)^{n_{\operatorname{RDO}}}(1 / 2)^{n_{\operatorname{BDD}}}\left(\frac{1}{4}\right)^{n_{\mathrm{BDD}}}
$$

Under the alternative hypothesis

$$
L=\left(\hat{z}_{0}\right)^{n_{\text {BDO }}}\left(\hat{z}_{1}\right)^{n_{\text {BBD }}}\left(\hat{z}_{2}\right)^{n_{\mathrm{BBD} 2}}
$$

## Testing for Linkage

- Evaluate likelihood at null hypothesis
- Evaluate likelihood at MLE
- Compare alternatives using likelihood ratio test


## Commonly Used Test Statistics

$$
\begin{aligned}
L O D & =\log _{10} \frac{L\left(\hat{z}_{0}, \hat{z}_{1}, \hat{z}_{2}\right)}{L\left(z_{0}=1 / 4, z_{1}=1 / 2, z_{2}=1 / 4\right)} \\
\chi^{2} & =2 \ln \frac{L\left(\hat{z}_{0}, \hat{z}_{1}, \hat{z}_{2}\right)}{L\left(z_{0}=1 / 4, z_{1}=1 / 2, z_{2}=1 / 4\right)} \\
& =2 \ln L\left(\hat{z}_{0}, \hat{z}_{1}, \hat{z}_{2}\right)-2 \ln L\left(z_{0}=1 / 4, z_{1}=1 / 2, z_{2}=1 / 4\right)
\end{aligned}
$$

## Example

$5 x$

$\operatorname{IBD}=1$ 1 2
$5 x \quad 1 / 2 \quad 1 \sqrt{2}$

$\operatorname{IBD}=2 \quad 2 \sqrt{2}$
$2 \sqrt{2}$

## Example

- Assume that 10 sib-pairs are examined
- 5 share 2 alleles IBD
- 5 share 1 allele IBD
- Calculate likelihood for null
- Calculate MLEs
- Calculate LOD score
- Evaluate LOD for each pair


## In real life...

- Markers are only partially informative
- IBD sharing is equivocal
- Some uncertainty removed by examining relatives
- Need an alternative likelihood
- Should allow for partially informative data


## Desirable Properties

- Also depends on parameters $\mathrm{z}_{0}, \mathrm{z}_{1}, \mathrm{z}_{2}$
- Probability of sharing 0,1 and 2 alleles IBD
- Can incorporate partial information on IBD sharing
- For fully informative data, equivalent to previous likelihood


## For A Single Family

$$
L_{i}=\sum_{j=0}^{2} P(I B D=j \mid A S P) P\left(\text { Genotypes }_{i} \mid I B D=j\right)=\sum_{j=0}^{2} z_{j} w_{i j}
$$

Risch (1990) defines

$$
w_{i j}=P\left(\text { Genotypes }_{i} \mid I B D=j\right)
$$

We only need proportionate $w_{i j}$

## Likelihood and LOD Score

$$
\begin{aligned}
& L\left(z_{0}, z_{1}, z_{2}\right)=\prod_{i} \sum_{j} z_{j} w_{i j} \\
& L O D=\log _{10} \prod_{1} \frac{\hat{y}_{0} w_{10}+\hat{z}_{1} w_{1}+\hat{z}_{2} w_{12}}{1 / w_{10}+1 / w_{11}+1 / 4 w_{12}}
\end{aligned}
$$

The MLS statistic is the LOD evaluated at the MLEs of $\mathrm{z}_{0}, \mathrm{z}_{1}, \mathrm{z}_{2}$

## Example: Scoring of $\mathrm{w}_{\mathrm{ij}}$



In this case, only one of the weights is non-zero for each family.

## More interesting examples: $\mathrm{w}_{\mathrm{ij}}$



In these cases, multiple weights are non-zero (but equal) for each family.

## More interesting examples: $\mathrm{w}_{\mathrm{ij}}$



In this case, relative weights depend on allele frequency.

## How to maximize likelihood?

- If all families are informative
- Use sample proportions of IBD=0, 1, 2
- If some families are uninformative
- Use an E-M algorithm
- At each stage generate complete dataset with fractional counts
- Iterate until estimates of LOD and z parameters are stable


## Assigning Partial Counts in E-M

$$
\begin{aligned}
P(I B D & =j \mid \text { Genotypes })= \\
& =\frac{P(I B D=j \mid A S P) P(\text { Genotypes } \mid I B D=j)}{L_{i}} \\
& =\frac{P(I B D=j \mid A S P) P(\text { Genotypes } \mid I B D=j)}{\sum_{k=0}^{2} P(I B D=k \mid A S P) P(\text { Genotypes } \mid I B D=k)} \\
& =\frac{Z_{j} w_{i j}}{\sum_{k=0}^{2} Z_{k} w_{i k}}
\end{aligned}
$$

## Example



Assume a bi-allelic marker where the two alleles have identical frequencies.

## Example of E-M Steps

| Parameters |  |  |  | Equivocal Families |  |  |  |  |  |  |  | Other |  |  |
| :---: | :---: | :---: | :---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: |
| z0 | z1 | z2 | IBD=0 | IBD=1 | IBD=2 | IBD=2 | LOD | LODi | LODu |  |  |  |  |  |
| 0.250 | 0.500 | 0.250 | 0.56 | 2.22 | 2.22 | 5 | 0.00 | 0.00 | 0.00 |  |  |  |  |  |
| 0.056 | 0.222 | 0.722 | 0.08 | 0.66 | 4.26 | 5 | 3.19 | 2.30 | 0.89 |  |  |  |  |  |
| 0.008 | 0.066 | 0.926 | 0.01 | 0.17 | 4.82 | 5 | 4.01 | 2.84 | 1.16 |  |  |  |  |  |
| 0.001 | 0.017 | 0.982 | 0.00 | 0.04 | 4.96 | 5 | 4.20 | 2.97 | 1.23 |  |  |  |  |  |
| 0.000 | 0.004 | 0.996 | 0.00 | 0.01 | 4.99 | 5 | 4.25 | 3.00 | 1.24 |  |  |  |  |  |
| 0.000 | 0.001 | 0.999 | 0.00 | 0.00 | 5.00 | 5 | 4.26 | 3.01 | 1.25 |  |  |  |  |  |
| 0.000 | 0.000 | 1.000 | 0.00 | 0.00 | 5.00 | 5 | 4.26 | 3.01 | 1.25 |  |  |  |  |  |

## Properties of Pair Analyses Explored by Risch

- Effect of marker informativeness
- Effect of adding relative genotypes
- Size of genetic effect
- Degree of relationship


## PIC:

Measure of Marker Informativeness

- Probability that alleles of parent can be distinguished in offspring
- Botstein et al, 1980.
- Markers that could track dominant alleles
- Probability that parent is heterozygous and informative in relation to spouse


## PIC - Definition

- In general:

$$
\text { PIC }=1-\sum_{i=1}^{n} p_{i}^{2}-\sum_{i=1}^{n} \sum_{j=i+1}^{n} 2\left(p_{i} p_{j}\right)^{2}
$$

For a equally frequent alleles

$$
\text { PIC }=\frac{a-1}{a}-\frac{a-1}{a^{3}}
$$

- PIC <= Heterozygosity


## Some PICs and Heterozygosities

| Alleles | PIC | $\mathbf{H}$ |
| ---: | ---: | ---: |
| 2 | 0.38 | 0.50 |
| 3 | 0.59 | 0.67 |
| 4 | 0.70 | 0.75 |
| 5 | 0.77 | 0.80 |
| 8 | 0.86 | 0.88 |
| 10 | 0.89 | 0.90 |
| 20 | 0.95 | 0.95 |

## Marker Informativeness

Proportion of LOD Retained


## Marker Informativeness Gene of Modest Effect ( $\lambda_{0}=3$ )

Expected LOD Score


## Marker Informativeness Gene of Larger Effect ( $\lambda_{0}=10$ )

Expected LOD Score


## Genotypes of Other Family Members

- Expected LOD score decreases
- by < 33\% if only sib-pairs are typed
- by $<60 \%$ for second degree relatives
- by $<70 \%$ for third degree relatives
- Genotyping effort decreases by
- by 50\% if only sib-pairs are typed
- by 60\% if only second degree relatives typed
- by 75\% if only third degree relatives typed


## Quick Comment on Literature

- Greenwood and Schork (2004) suggested that uninformative families could bias MLS
- However, their results use a poor estimate for MLEs
- If an E-M algorithm is used, there is no problem


## Today ...

- Describe a likelihood model based on IBD sharing for pairs of individuals
- Model accommodates partially informative families
- Maximum LOD score can be calculated using an E-M algorithm


## Recommended Reading

- Risch (1990)
- Linkage Strategies for Genetically Complex Traits. III. The Effect of Marker Polymorphism on Analysis of Affected Relative Pairs
- Am J Hum Genet 46:242-253
- Introduces MLS method for linkage analysis
- Still, one of the best methods for analysis pair data
- Evaluates different sampling strategies
- Results were later corrected by Risch (1992)


## Recommended Reading

- Risch (1992)
- Corrections to Linkage strategies for genetically complex traits. III. The effect of marker polymorphism on analysis of affected relative pairs.
- Am J Hum Genet 51:673-675
- Evaluates utility of parental genotype data

