# Modeling IBD for <br> Pairs of Relatives 

Biostatistics 666
Lecture 17

## Previously ...

- Linkage Analysis of Relative Pairs
- IBS Methods
${ }^{\circ}$ Compare observed and expected sharing
- IBD Methods
${ }^{-}$Account for frequency of shared alleles
- Provide estimates of IBD sharing at each locus


## IBS Linkage Test

$$
\chi_{2 d f}^{2}=\sum_{i} \frac{\left(N_{I B S=i}-E\left[N_{I B S=i}\right]\right)^{2}}{E\left(N_{I B S=i}\right)}
$$

- $E\left(N_{I B S=i}\right)$ depends on $N$ and allele frequencies
- Bishop and Williamson (1990)


## Likelihood for Sibpair Data

$L_{i} \propto \sum_{j=0}^{2} P(I B D=j \mid A S P) P($ Genotypes $\mid I B D=j) \propto \sum_{j=0}^{2} z_{j} w_{i j}$

Risch (1990) defines

$$
\begin{gathered}
w_{i j} \propto P\left(\text { Genotypes }_{i} \mid I B D=j\right) \\
z_{i}=P(I B D=i \mid \text { affected relative pair })
\end{gathered}
$$

## MLS Statistic of Risch (1990)

$L\left(z_{0}, z_{1}, z_{2}\right)=\prod_{i} \sum_{j} z_{j} w_{i j}$
$L O D=\log _{10} \prod_{i} \frac{\hat{z}_{0} w_{i 0}+\hat{z}_{1} w_{i 1}+\hat{z}_{2} w_{i 2}}{1 / 4 w_{i 0}+1 / 2 w_{i 1}+1 / 4 w_{i 2}}=\frac{\chi^{2}}{2 \ln 10}$

The MLS statistic is the LOD evaluated at the MLEs of $\mathrm{z}_{0}, \mathrm{z}_{1}, \mathrm{z}_{2}$
The $\hat{\mathrm{z}}_{0}, \hat{\mathrm{z}}_{1}, \hat{\mathrm{z}}_{2}$ can be estimated using an E-M algorithm

## Today ...

- Predicting IBD for affected relative pairs
- Modeling marginal effect of a single locus
${ }^{-}$Relative risk ratio $\left(\lambda_{R}\right)$

The Possible Triangle for Sibling Pairs

- Plausible IBD values for affected siblings
- Refinement of the model of Risch (1990)


## Single Locus Model

1. Allele frequencies

- For normal and susceptibility alleles

2. Penetrances

- Probability of disease for each genotype
- Useful in exploring behavior of linkage tests
- A simplification of reality
- Ignore effect of other loci and environment


## Penetrance

- $f_{i j}=P($ Affected $\mid G=i j)$
- Probability someone with genotype ij is affected
- Models the marginal effect of each locus


## Using Penetrances

- Allele frequency $p$
- Genotype penetrances $f_{11}, f_{12}, f_{22}$
- Probability of genotype given disease
${ }^{-} P(G=i j \mid D)=$
- Prevalence
- K =


## Pairs of Individuals

- A genetic model can predict probability of sampling different affected relative pairs
- We will consider some simple cases:
- Unrelated individuals
- Parent-offspring pairs
- Monozygotic twins
- What do the pairs above have in common?


## What we might expect ...

- Related individuals have similar genotypes
- For a genetic disease...
- Probability that two relatives are both affected must be greater or equal to the probability that two randomly sampled unrelated individuals are affected


## Relative Risk and Prevalence

- In relation to affected proband, define
- $\mathrm{K}_{\mathrm{R}}$ prevalence in relatives of type R
- $\lambda_{R}=K_{R} / K$ increase in risk for relatives of type $R$
- $\lambda_{R}$ is a measure of the overall effect of a locus
- Useful for predicting power of linkage studies


## Unrelated Individuals

- Probability of affected pair

$$
\begin{aligned}
P(a \text { and } b \text { affected }) & =\mathrm{P}(a \text { affected }) \mathrm{P}(b \text { affected }) \\
& =\mathrm{P}(\text { affected })^{2} \\
& =\left[p^{2} f_{11}+2 p(1-p) f_{12}+(1-p)^{2} f_{22}\right]^{2} \\
& =K^{2}
\end{aligned}
$$

- For any two related individuals, probability that both are affected should be greater


## Monozygotic Twins

- Probability of affected pair

$$
\begin{aligned}
P(M Z \text { pair affected }) & =\sum_{G} P(G) P(a \text { affected } \mid G) P(b \text { affected } \mid G) \\
& =p^{2} f_{11}^{2}+2 p(1-p) f_{12}^{2}+(1-p)^{2} f_{22}^{2} \\
& =K_{M Z} K \\
& =\lambda_{M Z} K K
\end{aligned}
$$

- $\lambda_{\mathrm{MZ}}$ will be greater than for any other relationship


## Probability for Genotype Pairs

Child

| Parent | $A_{1} A_{1}$ | $A_{1} A_{2}$ | $A_{2} A_{2}$ |  |
| :---: | :---: | :---: | :---: | :---: |
| $A_{1} A_{1}$ | $p_{1}{ }^{3}$ | $p_{1}{ }^{2} p_{2}$ | 0 | $p_{1}{ }^{2}$ |
| $A_{1} A_{2}$ | $p_{1}{ }^{2} p_{2}$ | $p_{1} p_{2}$ | $p_{1} p_{2}{ }^{2}$ | $2 p_{1} p_{2}$ |
| $A_{2} A_{2}$ | 0 | $p_{1} p_{2}{ }^{2}$ | $p_{2}{ }^{3}$ | $p_{2}{ }^{2}$ |
|  | $p_{1}{ }^{2}$ | $2 p_{1} p_{2}$ | $p_{2}{ }^{2}$ | $N$ pairs |
|  |  |  |  |  |

## Probability of Genotype Pairs and Being Affected

Child

| Parent | $A_{1} A_{1}$ | $A_{1} A_{2}$ | $A_{2} A_{2}$ |  |
| :---: | :---: | :---: | :---: | :---: |
| $A_{1} A_{1}$ | $p_{1}{ }^{3} f_{11}{ }^{2}$ | $p_{1}{ }^{2} p_{2} f_{12} f_{12} f_{11}$ | 0 |  |
| $A_{1} A_{2}$ | $p_{1}{ }^{2} p_{2} f_{11} f_{12}$ | $p_{1} p_{2} f_{12}{ }^{2}$ | $p_{1} p_{2}{ }^{2} f_{12} f_{22}$ |  |
| $A_{2} A_{2}$ | 0 | $p_{1} p_{2}{ }^{2} f_{12} f_{22}$ | $p_{2}{ }^{3} f_{22}{ }^{2}$ |  |
|  |  |  |  | N pairs |

## Parent Offspring Pairs

- Probability of Affected Pair

$$
\begin{aligned}
P & =P(\text { parent and child affected }) \\
& =\sum_{\mathrm{G}_{\mathrm{p}}} \sum_{\mathrm{G}_{\mathrm{o}}} P\left(G_{P}, G_{o}\right) f_{G_{p}} f_{G_{o}} \\
& =\sum_{i} \sum_{j} \sum_{k} P(i, j, k) f_{i j} f_{i k} \\
& =p^{3} f_{11}^{2}+(1-p)^{3} f_{22}^{2}+p(1-p) f_{12}^{2}+2 p^{2}(1-p) f_{11} f_{12}+2 p(1-p)^{2} f_{22} f_{12} \\
& =K K_{O} \\
& =\lambda_{o} K K
\end{aligned}
$$

- $\lambda$ will be lower for other unilineal relationships
- $\lambda_{0}$ will be between 1.0 and $\lambda_{\mathrm{Mz}}$


## Point of Situation

- Probabilities of affected pairs for
- Unrelated Individuals
- Monozygotic Twins
- Parent-Offspring Pairs
- Each of these shares a fixed number of alleles IBD ...


## For a single locus model...

$\lambda_{\text {IBD }=2}=\lambda_{M Z}$
$\lambda_{\text {IBD }=1}=\lambda_{O}$
$\lambda_{\text {IBD }=0}=1$
$K_{I B D=2}=K_{M Z}$
$K_{\text {IBD }=1}=K_{O}$
$K_{\text {IBD }=0}=1$

- Model ignores contribution of other genes and environment

Simple model that allows for useful predictions

- Risk to half-siblings
- Risk to cousins
- Risk to siblings


## Affected Half-Siblings

- IBD sharing
- 0 alleles with probability $50 \%$
- 1 allele with probability $50 \%$

This gives ...

$$
\begin{gathered}
\lambda_{H}=1 / 2 \lambda_{O}+1 / 2=1 / 2\left(\lambda_{O}+1\right) \\
K_{H}=1 / 2 K_{O}+1 / 2 K=1 / 2\left(K_{O}+K\right)
\end{gathered}
$$

## Uni-lineal Relationships

$$
\begin{gathered}
\lambda_{R}=P(I B D=1 \mid R) \lambda_{O}+P(I B D=0 \mid R) \\
K_{R}=P(I B D=1 \mid R) K_{O}+P(I B D=0 \mid R) K
\end{gathered}
$$

$P(I B D=1)$ decreases $50 \%$ with
increasing degree of relationship
( $\lambda_{R}-1$ ) also decreases $50 \%$ with increasing degree of unilineal relationship

## Affected Sibpairs

- IBD sharing ...
- 0 alleles with probability $25 \%$
- 1 alleles with probability $50 \%$
- 2 alleles with probability 25\%
- This gives ...
$\lambda_{S}=1 / 4 \lambda_{M Z}+1 / 2 \lambda_{O}+1 / 4=1 / 4\left(\lambda_{M Z}+2 \lambda_{O}+1\right)$
which implies
$\lambda_{M Z}=4 \lambda_{S}-2 \lambda_{O}-1$


## Examples: Full Penetrance

| Recessive |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| p | f11 | f12 | f22 | K | Lambdas |  |  |
|  |  |  |  |  | MZ | Offspring | Sibling |
| 0.001 | 0 | 0 | 1 | 0.000001 | 1000000 | 1000 | 250500 |
| 0.01 | 0 | 0 | 1 | 0.0001 | 10000 | 100 | 2550 |
| 0.1 | 0 | 0 | 1 | 0.01 | 100 | 10 | 30 |

Dominant

| $\mathbf{p}$ | $\mathbf{f 1 1}$ | $\mathbf{f 1 2}$ | $\mathbf{f 2 2}$ | K |  |  |  |  |  |  | Lambdas |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: |
| 0.001 | 0 | 1 | 1 | 0.002 | 500.25 | Offspring | Sibling |  |  |  |  |  |  |
| 0.01 | 0 | 1 | 1 | 0.02 | 50.25 | 25.50 | 250.56 |  |  |  |  |  |  |
| 0.1 | 0 | 1 | 1 | 0.19 | 5.26 | 3.02 | 3.08 |  |  |  |  |  |  |

## Examples: Incomplete Penetrance

Recessive

|  |  |  |  | Lambdas |  |  |  |  |
| ---: | :---: | :---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\mathbf{p}$ | f11 | f12 | f22 | K |  | MZ | Offspring | Sibling |
| 0.001 | 0.001 | 0.001 | 1 | 0.001 | 2.0 | 1.0 | 1.2 |  |
| 0.01 | 0.001 | 0.001 | 1 | 0.001 | 83.5 | 1.8 | 22.0 |  |
| 0.1 | 0.001 | 0.001 | 1 | 0.01 |  | 82.8 | 8.4 | 25.2 |

## Dominant

|  |  |  |  | Lambdas |  |  |  |  |
| ---: | :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| p | f11 | f12 | f22 | K |  | MZ | Offspring | Sibling |
| 0.001 | 0.001 | 1 | 1 | 0.003 | 223 | 112 | 112 |  |
| 0.01 | 0.001 | 1 | 1 | 0.02 | 46 | 23 | 23 |  |
| 0.1 | 0.001 | 1 | 1 | 0.19 | 5 | 3 | 3 |  |

## Examples: Small Effects

## Smaller Effects

|  |  |  |  |  | Lambdas |  |  |
| :---: | :---: | :---: | :---: | :---: | ---: | ---: | ---: | ---: |
| $\mathbf{p}$ | f11 | f12 | f22 | K | MZ | Offspring | Sibling |
| 0.1 | 0.01 | 0.02 | 0.04 | 0.012 | 1.2 | 1.1 | 1.1 |
| 0.1 | 0.01 | 0.08 | 0.16 | 0.024 | 2.6 | 1.8 | 1.8 |
| 0.1 | 0.02 | 0.16 | 0.32 | 0.048 | 2.6 | 1.8 | 1.8 |
| 0.2 | 0.01 | 0.02 | 0.04 | 0.014 | 1.2 | 1.1 | 1.1 |
| 0.2 | 0.01 | 0.08 | 0.16 | 0.038 | 2.1 | 1.6 | 1.6 |
| 0.2 | 0.02 | 0.16 | 0.32 | 0.08 | 2.1 | 1.6 | 1.6 |

## Multiple susceptibility loci...

- $\lambda$ are upper bound on effect size for one locus
- $\lambda$ decay rapidly for distant relatives
- If genes act multiplicatively, we can multiply marginal $\lambda$ together


## Another interpretation...

$\lambda_{\text {IBD }=2}=\lambda_{M Z}=\frac{P(\text { affected } \mid I B D=2 \text { with affected relative })}{P(\text { affected })}$

$$
\lambda_{I B D=1}=\lambda_{O}=\frac{P(\text { affected } \mid I B D=1 \text { with affected relative })}{P(\text { affected })}
$$

$$
\lambda_{I B D=0}=1=\frac{P(\text { affected } \mid I B D=0 \text { with affected relative })}{P(\text { affected })}
$$

## Bayes' Theorem: <br> Predicting IBD Sharing

$P(I B D=i \mid$ affected pair $)=$

$$
\begin{aligned}
& =\frac{P(I B D=i) P(\text { affected pair } \mid I B D=i)}{\sum_{j} P(I B D=j) P(\text { affected pair } \mid I B D=j)} \\
& =\frac{\lambda_{I B D=i}}{\sum_{j} P(I B D=j) \lambda_{I B D=i}}
\end{aligned}
$$

## Sibpairs

Expected Values for $z_{0}, z_{1}, z_{2}$

$$
\begin{aligned}
& \mathrm{z}_{0}=0.25 \frac{1}{\lambda_{s}} \\
& \mathrm{Z}_{1}=0.50 \frac{\lambda_{o}}{\lambda_{\mathrm{s}}} \\
& \mathrm{z}_{2}=0.25 \frac{\lambda_{\mathrm{MZ}}}{\lambda_{\mathrm{s}}}
\end{aligned}
$$

$1 \leq \lambda_{o} \leq \lambda_{s} \leq \lambda_{M Z}$ for any genetic model

## Maximum LOD Score (MLS)

- Powerful test for genetic linkage
- Likelihood model for IBD sharing
- Accommodates partially informative families
- MLEs for IBD sharing proportions
- Can be calculated using an E-M algorithm
- Shortcoming:
- Sharing estimates may be implausible


## Possible Triangle



## Possible Triangle



## Intuition

- Under the null
- True parameter values are ( $1 / 4,1 / 2,1 / 4$ )
- Estimates will wobble around this point
- Under the alternative
- True parameter values are within triangle
- Estimates will wobble around true point


## Idea (Holmans, 1993)

Testing for linkage

- Do IBD patterns suggest a gene is present?
- Focus on situations where IBD patterns are compatible with a genetic model
- Restrict maximization to possible triangle


## The possible triangle method

1. Estimate $\mathbf{z}_{0}, \mathbf{z}_{1}, \mathbf{z}_{2}$ without restrictions
2. If estimate of $z_{1}>1 / 2$ then $\ldots$
a) Repeat estimation with $z_{1}=1 / 2$
b) If this gives $z_{0}>1 / 4$ then revert to null (MLS=0)
3. If estimates imply $2 z_{0}>z_{1}$ then ...
a) Repeat estimation with $z_{1}=2 z_{0}$
b) If this gives $z_{0}>1 / 4$ then revert to null (MLS=0)
4. Otherwise, leave estimates unchanged.

## Possible Triangle



Holman's Example:

| IBD | Pairs |
| :--- | :--- |
| 0 | 8 |
| 1 | 60 |
| 2 | 32 |

MLS = 4.22 (overall)
MLE $=(0.08,0.60,0.32)$
MLS $=3.35$ (triangle)
MLE $=(0.10,0.50,0.40)$

## MLS Combined With Possible Triangle

- Under null, true $\mathbf{z}$ is a corner of the triangle
- Estimates will often lie outside triangle
- Restriction to the triangle decreases MLS
- MLS threshold for fixed type I error decreases
- Under alternative, true $\mathbf{z}$ is within triangle
- Estimates will lie outside triangle less often
- MLS decreases less
- Overall, power should be increased


## Example

- Type I error rate of 0.001
- LOD of 3.0 with unrestricted method
- Risch (1990)
- LOD of 2.3 with possible triangle constraint
- Holmans (1993)
- For some cases, almost doubles power


## Recommended Reading

- Holmans (1993)

Asymptotic Properties of
Affected-Sib-Pair Linkage Analysis
Am J Hum Genet 52:362-374

- Introduces possible triangle constraint
- Good review of MLS method


## Reference

- Risch (1990)

Linkage strategies for genetically complex traits. I. Multi-locus models. Am. J. Hum. Genet. 46:222-228

- Recurrence risks for relatives.
- Examines implications of multi-locus models.

