Introduction

There is a worldwide epidemic of obesity and type 2 diabetes across all age groups, especially in industrialized countries [1]. In the United States alone, over two-thirds of the population has a body mass index (BMI) of 25 kg/m² or greater and is thus overweight [2,3]. Being overweight is a well-established risk factor for many chronic diseases, such as type 2 diabetes, hypertension, and cardiovascular events [4], and increases in BMI are associated with higher all-cause mortality [5,6]. The economic cost attributable to obesity in the United States has been estimated to be as high as $100 billion/yr [7], and includes not only direct health care costs but also the cost of lost productivity in affected individuals [8].

Individual susceptibility to obesity is thought to be determined by interactions between an individual's genetic traits. PLoS Genet 3(7): e115. doi:10.1371/journal.pgen.0030115

The obesity epidemic is responsible for a substantial economic burden in developed countries and is a major risk factor for type 2 diabetes and cardiovascular disease. The disease is the result not only of several environmental risk factors, but also of genetic predisposition. To take advantage of recent advances in gene-mapping technology, we executed a genome-wide association scan to identify genetic variants associated with obesity-related quantitative traits in the genetically isolated population of Sardinia. Initial analysis suggested that several SNPs in the FTO and PFKP genes were associated with increased BMI, hip circumference, and weight. Within the FTO gene, rs9930506 showed the strongest association with BMI ($p = 8.6 \times 10^{-7}$), hip circumference ($p = 3.4 \times 10^{-8}$), and weight ($p = 9.1 \times 10^{-7}$). In Sardinia, homozygotes for the rare “G” allele of this SNP (minor allele frequency = 0.46) were 1.3 BMI units heavier than homozygotes for the common “A” allele. Within the PFKP gene, rs6602024 showed very strong association with BMI ($p = 4.9 \times 10^{-6}$). Homozygotes for the rare “A” allele of this SNP (minor allele frequency = 0.12) were 1.8 BMI units heavier than homozygotes for the common “G” allele. To replicate our findings, we genotyped these two SNPs in the GenNet study. In European Americans ($N = 1,496$) and in Hispanic Americans ($N = 839$), we replicated significant association between rs9930506 in the FTO gene and BMI ($p$-value for meta-analysis of European American and Hispanic American follow-up samples, $p = 0.001$), weight ($p = 0.001$), and hip circumference ($p = 0.0005$). We did not replicate association between rs6602024 and obesity-related traits in the GenNet sample, although we found that in European Americans, Hispanic Americans, and Afro-Caribbean Americans, homozygotes for the rare “A” allele were, on average, 1.0–3.0 BMI units heavier than homozygotes for the more common “G” allele. In summary, we have completed a whole-genome–association scan for three obesity-related quantitative traits and report that common genetic variants in the FTO gene are associated with substantial changes in BMI, hip circumference, and body weight. These changes could have a significant impact on the risk of obesity-related morbidity in the general population.


Abbreviations: AA, African American; BMI, body mass index; CEU, Utah residents with ancestry from northern and western Europe; EA, European American; FDR, false-discovery rate; HA, Hispanic American; LD, linkage disequilibrium; YRI, Yoruba in Ibadan, Nigeria

* To whom correspondence should be addressed. E-mail: schlessingerd@grc.nia.nih.gov (DS); goncalo@umich.edu (GRA)

† These authors contributed equally to this work.
and we genotyped an additional 1,412 individuals using the Affymetrix 10,000 SNP Mapping Array in other individuals in the family [23,24].

characterize the haplotypes in each stretch and impute them for participation in the study; a detailed account of the family structures we examined is available elsewhere [20]. We took advantage of the relatedness among individuals in our sample to substantially reduce study costs [23]. Specifically, because our sample includes many large families, we reasoned that genotyping a relatively small number of markers in all individuals would allow us to identify shared haplotype stretches within each family. We could then genotype a subset of the individuals in each family at higher density to characterize the haplotypes in each stretch and impute them in other individuals in the family [23,24].

For the analyses presented here, we genotyped 3,329 individuals using the Affymetrix 10,000 SNP Mapping Array and we genotyped an additional 1,412 individuals using the Affymetrix 500,000 SNP Mapping Array Set. The genotyped individuals were selected to represent the largest families in our sample, without respect to phenotype. The high-density arrays were generally used to genotype both parents and one child (in larger sibships) or just the parents (in smaller sibships); the lower density arrays were used to genotype everyone else. Except when parents and offspring were genotyped in the same family, we tried to ensure that individuals genotyped with the high-density array were only distantly related to one another. For the 2,893 individuals that were genotyped with the 10,000 SNP arrays only, we used a modified version of the Lander-Green algorithm [25,26] to probabilistically infer missing genotypes [24]. Our approach for estimating missing genotypes is implemented in MERLIN (http://www.sph.umich.edu/csg/abecasis/MERLIN/) and described in detail elsewhere [24].

We tested 362,129 SNPs for association with three obesity-related quantitative traits (BMI, hip circumference, and weight). Height was included as a covariate in analysis of hip circumference and weight. In addition, we included age and sex as covariates in every analysis. The genomic control parameter [29] for our initial analysis of each trait ranged from 1.07 to 1.09, indicating that our estimated test statistics might be slightly inflated. This is likely due to unaccounted-for distant relationships among the sampled individuals. All results presented in our tables have been adjusted using the method of genomic control [29]. After adjustment, we observed no significant excess of results exceeding liberal significance thresholds. For example, the proportion of test statistics that were significant at \( \alpha = 0.001 \) was 0.00098.

Results

We tested 362,129 SNPs for association with three obesity-related quantitative traits (BMI, hip circumference, and weight). Height was included as a covariate in analysis of hip circumference and weight. In addition, we included age and sex as covariates in every analysis. The genomic control parameter [29] for our initial analysis of each trait ranged from 1.07 to 1.09, indicating that our estimated test statistics might be slightly inflated. This is likely due to unaccounted-for distant relationships among the sampled individuals. All results presented in our tables have been adjusted using the method of genomic control [29]. After adjustment, we observed no significant excess of results exceeding liberal significance thresholds. For example, the proportion of test statistics that were significant at \( \alpha = 0.001 \) was 0.00098.

Results of our initial association analysis are summarized in Figure 1 and in Table 1. We used the false-discovery rate (FDR) to select a small set of very promising trait SNP associations for rapid replication. Using an FDR [30] of 20% highlighted a small set of SNPs for each trait. This set include the top eight SNP association results for hip circumference and weight (FDR = 0.013 and FDR = 0.16, respectively) and the top nine SNP association results for BMI (FDR = 0.20).

Eight of the SNPs listed in Table 1 overlap among the three traits. In particular, SNP rs9390506 and a cluster of nearby SNPs on Chromosome 16 show strong association with BMI (\( p = 8.6 \times 10^{-7} \)), hip circumference (\( p = 3.4 \times 10^{-4} \)) and weight (\( p = 9.1 \times 10^{-7} \)). Two of the associated SNPs in the cluster, rs9399609 and rs9926289, fall within an intronic region where sequence is strongly conserved across species. For comparative purposes, using a conservative Bonferroni correction aimed at an overall type I error rate of 0.05 (one false positive per 20 genome-scans), would result in a significance threshold of \( 1.4 \times 10^{-7} \).

This cluster of SNPs on Chromosome 16 overlaps the FTO [31] gene, an extremely large gene whose exons span \( >400 \text{ kb} \) (Figure 2). KIAA1005, a gene of unknown function, also maps nearby. The FTO gene has not been previously implicated in obesity, but it maps to a region where linkage to BMI has been
Figure 1. Log of p-Value for Single Marker Association Analysis with Three Obesity-Related Traits
Locations of PFKP and FTO genes are highlighted.
doi:10.1371/journal.pgen.0030115.g001
Table 1. Markers Showing Strongest Evidence for Association

<table>
<thead>
<tr>
<th>Trait</th>
<th>SNP</th>
<th>Chromosome</th>
<th>Position</th>
<th>Allele</th>
<th>Frequency</th>
<th>Effect (s.d.)</th>
<th>H²</th>
<th>p-Value</th>
<th>FDR</th>
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<tbody>
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<td>BMI</td>
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<td>52387965</td>
<td>A</td>
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<td>-0.132</td>
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<td>8.6E-07</td>
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<td>G</td>
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<td>-0.128</td>
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</tr>
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<td>-0.152</td>
<td>1.51%</td>
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<tr>
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<td>-0.152</td>
<td>1.51%</td>
<td>8.5E-06</td>
<td>0.006</td>
</tr>
<tr>
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<td>-0.152</td>
<td>1.51%</td>
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<td>Weight</td>
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<tr>
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<td>G</td>
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<td>0.10</td>
</tr>
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<td>85760188</td>
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<td>-0.105</td>
<td>0.96%</td>
<td>1.6E-05</td>
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</tr>
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</table>

All positions refer to the May 2004 genome assembly. The effect is measured in standard deviation units (so that an effect of -0.132 indicates that each additional copy of the allele decreases trait values by 0.132 standard deviations on average). The H² column indicates the proportion of the trait variance that could be explained by the SNP. The FDR column estimates the proportion of false positives incurred by declaring a particular SNP significant (so that an FDR of 0.09 indicates that declaring all SNPs with a smaller p-value significant is expected to result in about 9% false-positive findings).

s.d., standard deviation.
doi:10.1371/journal.pgen.0030115.t001

reported in two previous genome-wide linkage scans (LOD = 3.2 in the Framingham Heart Study [32] and LOD = 2.2 in the families with white ancestry from the Family Blood Pressure Program [33]). Furthermore, a syndrome that results from deletion of this region of Chromosome 16q includes obesity as one of its features [34].

Although multiple SNPs within FTO show evidence for association, these do not point to multiple independently associated SNPs—rather, it is likely they are all in disequilibrium with the same causal variant(s). In a sequential analysis in which we selected the best SNP for each trait and then conditioned on it to successively select the next best SNP, only one FTO SNP was selected (results presented in Table S2). This result is consistent with the fact that the SNPs fall in a region of strong linkage disequilibrium, both in Sardinia and in the HapMap (Figure 2B).

Our FDR analysis of BMI selected one additional SNP outside this cluster, rs6602024 (Figure 3). This SNP maps to Chromosome 10 and shows association with BMI (p = 4.9 × 10^{-6}, weight (1.6 × 10^{-5}), and hip circumference (p = 0.00047). The SNP maps to the platelet-type phosphofructokinase (PFKP) gene, which acts as a major rate-limiting enzyme in glycolysis, converting D-fructose-6-phosphate to fructose-1,6-bisphosphate [35]. Alterations in the structure or regulation of PFKP could alter the balance between glycolysis and glycogen production, ultimately leading to obesity.

Table 2 shows the phenotypic effects associated with each of the two SNPs in our sample. Because rs9930506 is more common, it shows more significant association despite being associated with smaller phenotypic effects (the two homozygotes differ, on average, by ~1.5 BMI units). A rarer polymorphism, such as rs6602024, impacts only a smaller proportion of the population and shows less significant association, despite a larger difference between homozygotes means (which differ, on average, by ~2.9 BMI units). In each case, a more accurate estimate of the effect is provided by the regression model with age, sex, and (where appropriate) height as covariates. In a study, such as ours, that estimates effect sizes for many SNPs, statistical fluctuation means that some estimates will be slightly high and others will be low. SNPs that reach statistical significance are likely to include those for which effect size estimates are inflated (this is the winner’s curse phenomenon) [36], and thus we proceeded to replicate our top association signals in additional large samples.

To further investigate the association between rs9930506 and rs6602024 and obesity-related traits, we genotyped these SNPs in the GenNet study [37]. The study includes a series of families recruited through probands with elevated blood
The families included in this analysis comprise 3,467 individuals in total (1,101 African Americans [AA] in 369 families, 839 Hispanic Americans [HA] in 223 families, and 1,496 European Americans [EA] in 457 families). Overall, individuals in GenNet are heavier than those in our original Sardinian sample. Nevertheless, our findings strongly confirm evidence for association between rs9930506 and the three BMI-related traits (weight, hip circumference, and BMI). Specifically, rs9930506 showed association with all three traits among EA and HA in the GenNet study (meta-analysis of the EA and HA samples results in a \( p \)-value between 0.0005 and 0.001, depending on trait; see Table 3). The association is significant and in the same direction as in our original sample. The allele frequencies are also similar in all three samples, with a frequency of 0.46 in our Sardinian sample for allele “G” of rs9930506 and of 0.44 and 0.33 in the GenNet EA and HA samples, respectively. In the GenNet sample, homozygotes for the two rs9930506 alleles differ in weight by \( \sim 1.0 \text{ BMI units on average.} \)

We also examined the relationship between rs9930506 and the three traits in AA, but did not observe evidence for association within that group. In AA, allele “G” of marker rs9930506 has a somewhat lower frequency of 0.21. In addition, AA show quite distinct patterns of linkage disequilibrium (LD) and thus it is not surprising that the association does not replicate. For example, in the HapMap CEU sample the eight SNPs that show association with obesity-related traits in our sample are strongly associated

---

**Figure 2.** Association Results and LD Patterns in Region Surrounding the FTO Gene

(A) Summary of the association between SNPs in the region and BMI. The SNP showing strongest association (rs9930506) is highlighted. Other SNPs are colored according to their degree of disequilibrium with rs9930506 ranging from high (red), to intermediate (green), to low (blue). Transcripts are indicated at the bottom of the graph, with an arrow indicating transcript direction.

(B) Summary of the patterns of disequilibrium in the region in Sardinia and in two of the HapMap populations (CEU and YRI) [55]. The grey bar marks the region of association and facilitates comparisons between the panels.

doi:10.1371/journal.pgen.0030115.g002
with each other and tag a total of 38 different variants \((r^2 > 0.8)\). In contrast, in the HapMap Yoruba in Ibadan, Nigeria (YRI) the strength of LD in the region is greatly reduced such that rs9930506 is not in strong LD \((r^2, 0.3)\) with any of the other Chromosome 16 SNPs that show association in Sardinia.

In an attempt to fine-map association in the region, we decided to genotype the region of strong association in greater detail. In general, the study of samples from AA participants can afford an opportunity to fine-map association signals and even facilitate identification of the causal variants [38]. As noted above, a total of 38 different variants are in LD \((r^2 > 0.8)\), HapMap Utah residents with ancestry from northern and western Europe [CEU]) with the eight SNPs that are associated with obesity-related traits in our Sardinian sample. We selected an additional seven SNPs in the region to tag these 38 variants in samples with reduced LD. Together with rs9930506, these seven variants capture the other 30 SNPs with \(r^2 > 0.58\) (average \(r^2 = 0.87\), HapMap YRI). The results are summarized in Table 4 and show that, whereas all the variants show association in EA and HA, none of the variants shows association in AA. One possible explanation is that obesity in AA has a different genetic architecture. Alternatively, it is possible that because some of the variants are quite common in EA and HA but rare in AA, much larger sample sizes will be required to adequately gauge their effects (for example, rs1421085 and rs3751812 have minor allele frequencies >0.25 in these first two populations, but <0.11 in AA).

In contrast to rs9930506, we did not replicate association between SNP rs6602024 in the PFKP gene and the three obesity-related traits. The “A” allele was rare in all

---

**Figure 3.** Association Results and LD Patterns in Region Surrounding the **PFKP** Gene

(A) Summary of the association between SNPs in the region and BMI. The SNP showing strongest association (rs6602024) is highlighted. Other SNPs are colored according to their degree of disequilibrium with rs6602024, ranging from high (orange) to low (blue). Transcripts are indicated at the bottom of the graph, with an arrow indicating transcript direction.

(B) Summary of the patterns of disequilibrium in the region in Sardinia and in two of the HapMap populations (CEU and YRI) [55]. The grey bar marks the region of association and facilitates comparisons between the panels.

doi:10.1371/journal.pgen.0030115.g003
populations, with a frequency of 0.12 in our Sardinian sample, 0.11 in the HA and EA GenNet subsamples and 0.25 in the AA GenNet subsample. The results are summarized in Table 3 and show that, although homozygotes for the rare “A” allele at rs6602024 were on average heavier by ~1.0–3.0 BMI units than homozygotes for the “G” allele at the SNP, these homozygotes were rare and, overall, there was no significant association. Corroborating evidence that PFKP and rs6602024 are associated with BMI is the observation that a region of ~120 kb including the Pfkp gene has been implicated in a mouse model of obesity [39] (see Discussion). A definite assessment of the impact of PFKP on obesity-related quantitative traits in human populations will likely require examination of much larger sample sizes.

Our genotyping results also hint at the possible importance in Sardinia of other genes previously investigated as candidates influencing obesity and related traits (Tables S3–S5). When we evaluated evidence for association across previously identified candidate genes, we observed a small excess of nominally significant \( p \)-values. We tested 837 candidate SNPs in 74 candidate genes against three traits and found that 145 tests were significant at \( p < 0.05 \), corresponding to 5.8% of the 2,511 tests. We observed no such excess when the whole genome was considered.) Among the interesting candidates that show association in our sample are the two adiponectin receptor genes [40] ADIPOR1 (best single SNP \( p \)-value = 0.013, 0.027, and 0.016 for BMI, hip circumference, and weight) and ADIPOR2 (best \( p \)-values = 0.018, 0.019, 0.013) and the lipoprotein lipase gene, LPL [41] (best \( p \)-values = 0.014, 0.006, 0.018). Nevertheless, all the association signals observed in any of these previous candidate genes are far less significant than those in FTO or PFKP.

**Discussion**

FTO association provides an example of how genome-wide association studies can point to previously unsuspected candidate genes. An interstitial deletion overlapping the region produces human syndromic obesity [34] and a hint that the gene might be involved in stress responses stems from the observation that it is down-regulated when the heat

### Table 2. Effects Associated with the rs9930506 and rs6602024 SNPs

<table>
<thead>
<tr>
<th>Gene/SNP</th>
<th>Genotype Class</th>
<th>Frequency</th>
<th>Genotype Means</th>
</tr>
</thead>
<tbody>
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<td></td>
<td></td>
<td></td>
<td>Weight (kg) Hip (cm) BMI (kg/m²)</td>
</tr>
<tr>
<td><strong>FTO/rs9930506</strong></td>
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<td></td>
</tr>
<tr>
<td>G/G</td>
<td>0.21</td>
<td>69.8</td>
<td>101.0</td>
</tr>
<tr>
<td>G/A</td>
<td>0.50</td>
<td>67.3</td>
<td>99.7</td>
</tr>
<tr>
<td>A/G</td>
<td>0.29</td>
<td>66.7</td>
<td>98.3</td>
</tr>
<tr>
<td><strong>PFKP/rs6602024</strong></td>
<td>Estimated additive effect</td>
<td>0.46 (G)</td>
<td>1.636</td>
</tr>
<tr>
<td>A/A</td>
<td>0.01</td>
<td>75.6</td>
<td>103.2</td>
</tr>
<tr>
<td>A/G</td>
<td>0.21</td>
<td>68.0</td>
<td>100.2</td>
</tr>
<tr>
<td>G/G</td>
<td>0.77</td>
<td>67.2</td>
<td>99.2</td>
</tr>
<tr>
<td><strong>Estimated additive effect</strong></td>
<td>0.12 (A)</td>
<td>2.194</td>
<td>1.296</td>
</tr>
</tbody>
</table>

*Effect was estimated in a variance component model including sex, age, age², and (except for BMI) height as covariates.

**Table 3. Replication of Association between rs9930506/FTO/G Allele and Obesity-Related Traits**

<table>
<thead>
<tr>
<th>Sample</th>
<th>Average BMI</th>
<th>BMI Associated Effect (s.d.)</th>
<th>p-Value</th>
<th>Weight Associated Effect (s.d.)</th>
<th>p-Value</th>
<th>Hip Associated Effect (s.d.)</th>
<th>p-Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sardinia</td>
<td>26.4(4.1)</td>
<td>26.9(4.5)</td>
<td>27.9(5.1)</td>
<td>+0.132</td>
<td>0.000000086</td>
<td>+0.118</td>
<td>0.00000091</td>
</tr>
<tr>
<td>EA*</td>
<td>28.9(6.6)</td>
<td>28.6(6.2)</td>
<td>30.5(7.0)</td>
<td>+0.080</td>
<td>0.03</td>
<td>+0.071</td>
<td>0.03</td>
</tr>
<tr>
<td>HA*</td>
<td>28.4(6.0)</td>
<td>28.8(6.3)</td>
<td>29.7(7.1)</td>
<td>+0.122</td>
<td>0.03</td>
<td>+0.104</td>
<td>0.03</td>
</tr>
<tr>
<td>AA*</td>
<td>30.1(8.1)</td>
<td>30.8(8.5)</td>
<td>30.4(8.7)</td>
<td>-0.004</td>
<td>0.9</td>
<td>-0.006</td>
<td>0.5</td>
</tr>
<tr>
<td>Combined GenNet EA and HA (but excluding Sardinia)</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>0.001</td>
<td>—</td>
<td>0.001</td>
<td>—</td>
</tr>
<tr>
<td>Combined GenNet EA, HA, and AA (but excluding Sardinia)</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>0.007</td>
<td>—</td>
<td>0.015</td>
<td>—</td>
</tr>
</tbody>
</table>

The table summarizes the effect of allele “G” for rs9930506 in the original sample and in each of the replication samples examined. The allele has a frequency of 0.46 in our Sardinian sample and of 0.44, 0.33, and 0.21 in the EA, HA, and AA GenNet subsamples, respectively. Average BMI and standard deviation (inside parenthesis) is reported for each genotype class. The effect size (in standard deviation units, after normalizing the data and adjusting for covariates) and significance of the observed association is reported for each sample. Note that, although the effect observed in Sardinia is slightly larger than in the replication samples when measured in standard deviation units, the replication samples also appear to show greater variability in trait values overall (as seen by the larger standard deviations within each genotype class). The combined \( p \)-values are one-sided and test for an effect that is in the same direction as that seen in Sardinia.

*GenNet Family Blood Pressure Study.

s.d., standard deviation.

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shock response transcription factor *Hif1* is inhibited [42]. Because the gene has no recognizable functional domains and has not been studied in detail in experimental models, no putative function can be currently imputed that involves its action. The fact that *FTO* is associated not only with BMI but also with hip circumference and weight is consistent with previous analyses of heritability in our cohort [20]. The analyses suggested that 80% of the genetic variance of these traits is determined by common loci (individually, the traits have heritabilities between $\sim$30%–45%). Although the three traits examined here are correlated (all pairwise correlations consistently heavier than individuals with a *Combined GenNet EA, HA* allele are consistently heavier than individuals with a “G/G” genotype. There are of 72 such homozygotes in the replication sample (nine EA, ten HA, and 53 AA).

Table 5. Replication of Association between rs6602024/PFKP/A Allele and Obesity-Related Traits

<table>
<thead>
<tr>
<th>Sample</th>
<th>Average BMI</th>
<th>G/G</th>
<th>A/G</th>
<th>A/A</th>
<th>BMI Effect</th>
<th>p-Value</th>
<th>Weight Effect</th>
<th>p-Value</th>
<th>Hip Effect</th>
<th>p-Value</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sardinia</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.016</td>
<td>0.000049</td>
<td>0.167</td>
<td>0.00016</td>
<td>0.113</td>
<td>0.00047</td>
</tr>
<tr>
<td>EA*</td>
<td>28.9 (6.4)</td>
<td>29.5 (6.9)</td>
<td>32.2 (6.8)</td>
<td>0.071</td>
<td>0.3</td>
<td>0.075</td>
<td>0.2</td>
<td>0.083</td>
<td>0.2</td>
<td></td>
</tr>
<tr>
<td>HA*</td>
<td>28.8 (6.4)</td>
<td>28.4 (6.0)</td>
<td>31.2 (7.8)</td>
<td>0.000</td>
<td>0.9</td>
<td>0.008</td>
<td>1</td>
<td>0.034</td>
<td>0.6</td>
<td></td>
</tr>
<tr>
<td>AA*</td>
<td>29.9 (8.3)</td>
<td>30.9 (8.4)</td>
<td>30.9 (8.4)</td>
<td>0.055</td>
<td>0.3</td>
<td>0.042</td>
<td>0.4</td>
<td>0.071</td>
<td>0.2</td>
<td></td>
</tr>
<tr>
<td>Combined GenNet EA and HA</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>0.225</td>
<td></td>
<td>—</td>
<td>0.152</td>
<td>—</td>
<td>0.238</td>
</tr>
<tr>
<td>(but excluding Sardinia)</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>0.094</td>
<td></td>
<td>0.094</td>
<td></td>
<td>0.098</td>
<td></td>
</tr>
</tbody>
</table>

The table is analogous to Table 3, but focuses on allele “A” for the rs6602024 SNP. The allele has a frequency of 0.12 in our Sardinian sample, 0.11 in the EA and HA GenNet subsamples, and 0.25 in the AA GenNet subsample. We did not find a significant additive effect for this allele in the replication samples. However, note that homozygotes for the “A” allele are consistently heavier than individuals with a “G/G” genotype. There are of 72 such homozygotes in the replication sample (nine EA, ten HA, and 53 AA).

*GenNet Family Blood Pressure Study.*

doi:10.1371/journal.pgen.0030115.s005
of the forms has been previously implicated in obesity in humans, although PFKP is mutated in some cases of impaired glycogen synthesis (glycogen storage disease VII; see Online Mendelian Inheritance in Man, http://www.ncbi.nlm.nih.gov/entrez/dispomim.cgi?id=232800) [35]. It is of considerable interest that compared to the other isozymes, PFKP has lower affinity for fructose-6-phosphate and decreased inhibition by ATP [43]. Consequently, PFKP is the most stringently regulated, responding to small changes at typical metabolic levels of effectors [44]. Genetic variants in the enzyme could thus adjust the rate of glycolysis, shifting the balance of metabolism between gluconeogenesis and glucose assimilation—a possible step in the etiology of obesity. Additionally, it is intriguing that in mice a locus associated with obesity has been mapped to a 127-kb interval that includes Pkp3 [39]. The mouse locus shows strong evidence of interaction with diet, with different effects in mice fed high-fat and low-fat diets. One possibility is that greater homogeneity of diet in Sardinia facilitated mapping, but made replication in other populations more difficult.

How significant are the associations observed? The replication of the FTO association in two different populations indicates that it is likely to be important not only in Sardinia, but in many different populations. In contrast, the failure to replicate the PPKP association in other populations suggests that (a) the association we identified may refer to rarer, population-specific variants; (b) the effects of the locus may depend on genetic or environmental background; or (c) the association identified in our original sample is due to the statistical fluctuations inherent in testing hundreds of thousands of SNPs. As for the public health impact of the observed associations, a 1-unit increment in BMI has been associated with an 8% increase in the risk of coronary heart disease [45] and excess weight in middle life is associated with increased overall risk of death [46]. Thus, the alleles reported here, which shift BMI by 1–1.5 units, have effects that are not only statistically significant but could also have important health consequences. Furthermore, apart from the direct contribution of these gene variants, they provide an entree to the analysis of genes and pathways that contribute additionally, and open new routes to possible eventual intervention.

Note: After completing this manuscript, we became aware of additional evidence that supports our report of association between FTO and obesity-related traits. First, genotyping of 1,780 individuals from the SUVIMAX study [47,48] replicated association of allele rs9939506 with increased BMI (p = 0.006). Combined evidence from SUVIMAX, GenNet EA, and GenNet HA resulted in a replication p-value of 1.5 × 10⁻⁵. In addition, two other large independent studies also show association of SNPs in FTO with increased BMI [49,50]. Genotyping of the SUVIMAX sample did not provide evidence for association between rs6060204 and BMI.

Materials and Methods

Study sample. We recruited and phenotyped 6,148 individuals, male and female, ages 14–102 y, from a cluster of four towns in the Lanuza Valley [29]. During physical examination of each individual, a blood sample was collected (for DNA extraction) and anthropometric traits were recorded. Here, we report analyses of hip circumference, weight, and the derived quantity BMI (which is calculated from a combination of height and weight). Genotyping was carried out using the Affymetrix 10K and 500K chips (http://www.affymetrix.com) using standard protocols. Summary assessments of genotype data quality are provided in the Results section and in Table S1.

To follow up on SNPs rs9939506 and rs6060204, we genotyped and examined the association between these two SNPs and BMI, hip circumference, and body weight in the GenNet study. The study comprised 1,647 individuals, total 8,676 recruited between 1995 and 2004 (1,128 AA, 839 HA, and 1,496 EA). Individuals were recruited at two field centers: EA were recruited from Tecumseh, Michigan, and AA and HA were recruited from Maywood, Illinois. Participants were recruited from families starting from a proband with high blood pressure. DNA was available for 3,905 individuals (968 AA, 824 HA, and 1,471 EA). SNP genotyping was performed using the 5'-nucleotide-based assay (TaqMan; ABI, http://www.appliedbiosystems.com) analyzed on an ABI Prism 7900 Real Time PCR System. Within each ethnic group, genotype completeness rates exceed 98% and there was no evidence for deviation from Hardy–Weinberg equilibrium (p > 0.05).

Statistical analysis. To ensure adequate control of type I error rates, we applied an inverse normal transformation to each trait prior to analysis [29]. The inverse normal transformation reduces the impact of outliers and deviations from normality on statistical analysis. The transformation involves ranking all available phenotypes, transforming these ranks into quantiles and, finally, converting the resulting quantiles into normal deviates. We included sex, age, and age² as covariates in all analysis. Height was significantly associated with weight and hip circumference and was included as an additional covariate in analysis of those traits. We fitted a simple regression model to each trait and used a variance component approach to account for correlation between different observed phenotypes within each family. For individuals who had genotype data available, we coded genotypes as 0, 1, and 2 (depending on the number of copies of the risk allele being tested). For individuals with missing genotype data, we used the Lander–Green algorithm to estimate an expected genotype score (between 0 and 2) for each individual [24]. Briefly, to estimate each genotype score we first calculate the likelihood of the observed genotype data. Then, we instantiate each missing genotype to a specific value that maximizes the pedigree likelihood. The ratio of the two likelihoods gives a posterior probability that the instantiated genotype is true, conditional on all available data. Due to computational constraints, we divided large pedigrees into subunits with "bit-complexity" of 19 or less (typically, 20–25 individuals) before estimating missing genotypes.

Our analytical approach considers all observed or estimated genotypes (rather than focusing on alleles transmitted from heterozygous parents) and thus is not immune to effects of population stratification. In homogenous populations, this type of analysis is expected to be more powerful [51,52]. To adjust for the effects of population structure and cryptic relatedness among sampled individuals, we used the genomic control method to adjust our test statistics for each trait separately [29]. FDRs were calculated with R’s p.adjust() procedure using the method of Benjamini and Hochberg [30]. Since we had an initial failure to identify the strongest SNP as the SNP that showed strongest evidence for association, we also carried out a sequential stepwise analysis. In this analysis, we selected the best SNP for each trait, and then conditioned on it to successively select the next best SNP. This sequential analysis can help identify regions with multiple independent association signals. The stepwise analysis was repeated for five rounds.

Candidate gene analysis. We selected 74 candidate genes previously tested for association with obesity in humans [53]. For each gene, we first evaluated the ability of the Affymetrix SNPs to tag SNPs (MAF > 0.05) within 5–50 kb of the gene. For each gene, we report coverage statistics as well as neighboring Affymetrix CEU database [54]. We then evaluated evidence for association using all Affymetrix SNPs within each gene as well as neighboring Affymetrix SNPs that could be used to improve coverage (r² > 0.5). For each gene, we report coverage statistics as well as the SNP that showed strongest evidence for association.

We selected 74 genes that were previously targeted in associations studies aiming to identify genetic determinants of obesity in humans [53]: ACR, ACTN, ADIPOQ, ADIPOR1, ADIPOR2, ADRB1, ADRB2, ADD1, AGT, AKT1, ANG, ASPN, ATG1A, AR, BDNF, CASQ1, COL1A1, COMT, CRP, CYP11B2, DIO1, EFNA3, EXN, FOXC2, GAD2, GFTPT1, GHHR, GNAS, GB3, GPR10, H6PD, HSD11B1, HTR2C, ICAM1, IGF1, IGF2, ILS, IL6R, KCNJ11, KI, LEP, LEPR, LRPC, LPL, LTA, MCA4, MCHR1, MKKS, MTHFR, MTNP, NMB, NOS3, NPY, NPYR1, NR0B2, NTRK2, PAR6AB, PLIN, PPARC, PPARGC1A, PRDM2, PTEN, QPRT, REN, SCNN5A, SNAP25, USF1, VDR, UCP1, UCP2, UCP3, and VDR. We did not consider genes associated with drug-induced body weight gain or mitochondrial genes [53].

The following genes have previously been investigated for their
role in obesity and related traits but are not well tagged by SNPs in the Affymetrix array: ADRB3, DRD4, INS, and APOE.

### Supporting Information

**Table S1. Genotype Data for Sardinian Cohort**

Found at doi:10.1371/journal.pgen.0030115.s001 (47 KB DOC).

**Table S2. Results of Stepwise Analysis to Identify Independent Risk Alleles**

To generate this table, we first sought the most significantly associated allele in the genome. We then added this allele to our baseline model and repeated our genome scan to identify the next associated SNPs.

Found at doi:10.1371/journal.pgen.0030115.s002 (55 KB DOC).

**Table S3. Tag SNP That Shows Strongest Association with BMI for Each Previously Identified Candidate Gene**

The first column indicates the name of a previously identified candidate. The second column indicates the number of SNPs in our Affymetrix arrays that are either in the gene or constitute the best available tag ($r^2 > 0.5$) for a genic SNP. The next column indicates the number of HapMap SNPs within $±$ 5 kb of the gene and the proportion of these that are covered at $r^2 > 0.50$ or $r^2 > 0.80$. The next columns indicate the SNP that showed strongest association in our analysis, the $p$-value, the tested allele and its frequency, and the estimated additive effect. The last column corresponds to the FDR incurred when all tested SNPs are considered and this test is declared significant.

Found at doi:10.1371/journal.pgen.0030115.s003 (162 KB DOC).

**Table S4. Tag SNP That Shows Strongest Association with Hip Circumference for Each Previously Identified Candidate Gene**

The first column indicates the name of a previously identified candidate. The second column indicates the number of SNPs in our Affymetrix arrays that are either in the gene or constitute the best available tag ($r^2 > 0.5$) for a genic SNP. The next column indicates the number of HapMap SNPs within $±$ 5 kb of the gene and the proportion of these that are covered at $r^2 > 0.50$ or $r^2 > 0.80$. The next columns indicate the SNP that showed strongest association in our analysis, the $p$-value, the tested allele and its frequency, and the estimated additive effect. The last column corresponds to the FDR incurred when all tested SNPs are considered and this test is declared significant.

Found at doi:10.1371/journal.pgen.0030115.s004 (165 KB DOC).

**Table S5. Tag SNP That Shows Strongest Association with Weight for Each Previously Identified Candidate Gene**

The first column indicates the name of a previously identified candidate. The second column indicates the number of SNPs in our Affymetrix arrays that are either in the gene or constitute the best available tag ($r^2 > 0.5$) for a genic SNP. The next column indicates the number of HapMap SNPs within $±$ 5 kb of the gene and the proportion of these that are covered at $r^2 > 0.50$ or $r^2 > 0.80$. The next columns indicate the SNP that showed strongest association in our analysis, the $p$-value, the tested allele and its frequency, and the estimated additive effect. The last column corresponds to the FDR incurred when all tested SNPs are considered and this test is declared significant.

Found at doi:10.1371/journal.pgen.0030115.s005 (163 KB DOC).

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**IRB Approval.** The study, including the protocols for subject recruitment and assessment, the informed consent for participants (and Assent Forms for those 14–18 y old), and the overall analysis plan were reviewed and approved by IRB boards for the Istituto di Neurogenetica e Neurofarmacologia (INN; Cagliari, Italy), the MedStar Research Institute (responsible for intramural research at the National Institutes of Aging, Baltimore, Maryland, United States of America), and for the University of Michigan (Ann Arbor, Michigan, United States of America).


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**Competing interests.** The authors have declared that no competing interests exist.

### References